

# Probability and Random Variables

Introduction:- The primary goals of the subject to study about the principles of random signals and to provide tools to deal with systems involving random signals.

## Random Signal:-

The outcome of an experiment which cannot be predicted from its past values is known as random signal (or) non-deterministic signals.

Eg:- Examples of Random Signals

- i) Now considering experiment of <sup>tossing a coin</sup> the outcome should be either Head (or) Tail.
- ii) while listening to an ordinary radio receiver sum noise will occur, that noise is called "hiss" noise.
- iii) In a Television System noise shows up in the form of picture interference, that noise called snow.

## Deterministic Signals:-

The outcome of an experiment to be predicted from its positive values is known as deterministic signals.

## Examples of Deterministic Signal:-

- i) Consider an experiment to study about the "where the sun rises".
- ii) The outcome of above experiment to be conducted on any day is the same that is sun rises in east.

## Probability Introduced through sets and relative frequency.

A Set is a collection of objects, the objects are also called elements of the Set. A Set is denoted by a capital letter and the elements are denoted by lower case letters.

ex:  $a \in A$

\* Any set is represented by two braces

$\{ \}$

## Countable Set :-

A Set is said to be countable if the elements can be put in one to one correspondence with natural numbers.

ex: A set of natural no's

$$N = \{1, 2, 3, 4\}$$

A set of odd no's

Uncountable Set : A set is not countable

then it is called uncountable set

ex:  $C = \{0.5 \leq x \leq 8.5\}$

Empty set & null set : A set is said to be empty if it has no elements then set is called null set

\*  $\emptyset$  is denoted by the symbol  $\emptyset$ .

ex:  $P = \{0, 0\}$  is not a null set.

Finite set : A finite set is one that is either empty or has elements.

ex:  $D = \{0, 0\}$  is a finite set.

$E = \{2, 4, 6, 8, 10, 12, 14\}$  is a finite set.

Infinite set : If a set is not a finite then

is called infinite set  $\{1, 2, 3, 4, \dots\}$

Countable infinite set :

An infinite set having a countable elements is called countable infinite set

ex: The set having "N" no. of natural numbers  $N = \{1, 2, 3, 4, \dots, \infty\}$

Uncountable infinite set :

An infinite set having uncountable elements is called uncountable infinite set.

ex:  $F = \{-50 \leq x \leq 90.0\}$

Subset : A set "x" is a subset of "y"

element of set "y".

\* It is denoted by a " $x \leq y$ ".

Equal Sets :-

A two sets "x" and "y" are said to be equal if both the sets are contained all common elements.

\* It is denoted by  $x = y$  &  $x \leq y$  and  $y \leq x$

Ex:  $x = \{1, 3, 5, 7\}$ ,  $y = \{1, 3, 5, 7\}$

Disjoint Sets :-

If two sets "x" and "y" are said to be disjoint sets if they contain no common elements.

Ex:  $x = \{1, 3, 5, 7\}$ ,  $y = \{2, 4, 6, 8\}$

Universal Set :-

A set which contains all the elements of a given system is called universal set.

\* It is denoted by "S".

Ex: - Rolling a die,  $\{1, 2, 3, 4, 5, 6\}$

Universal set =  $\{1, 2, 3, 4, 5, 6\}$

Relative Frequency :-

If we are conducting a particular experiment more number of times by just

observing the output is called Relative Frequency.

\* let "A" be a particular event in this sample space "S" then the probability of event "A" is denoted by  $P(A)$ .

\* Suppose the Random Experiment is repeated "n" number of times if the event "A" occurs  $n(A)$  times then the probability of event "A" is denoted by relative frequency of event "A" when  $n \rightarrow \infty$

$\therefore$  The probability of event "A" is defined as

$$P(A) = \lim_{n \rightarrow \infty} \frac{n(A)}{n}$$

where  $\frac{n(A)}{n}$  is called Relative Frequency of A

$n(A)$  means how many times a particular event occurs.

"n" means how many time the Experiment is repeated.

1. A die is thrown 1000 times with frequencies for the outcomes 1-6 given by

Outcomes	1	2	3	4	5	6
Frequencies	179	150	157	149	195	190

Find the probability of outcome "1" for the experiment is repeated 1000 times.

Sol: (i) P {outcome "1" for the experiment is repeated 1000 times} =  $\lim_{n \rightarrow \infty} \frac{n(1)}{n}$

$$= \lim_{n \rightarrow \infty} \frac{179}{1000} = \frac{179}{1000} = 0.179$$

(ii) P {outcome "2" for the experiment is repeated 1000 times} =  $\lim_{n \rightarrow \infty} \frac{150}{1000}$

$$= \frac{150}{1000} = 0.15$$

(iii) P {outcomes "3" for the experiment is repeated 1000 times} =  $\lim_{n \rightarrow \infty} \frac{157}{1000}$

$$= \frac{157}{1000} = 0.157$$

(iv) P {outcomes "4" for the experiment repeated is 1000 times} =  $\lim_{n \rightarrow \infty} \frac{149}{1000}$

$$= \frac{149}{1000} = 0.149$$

(v) P {outcomes "5" for the experiment repeated is 1000 times} =  $\lim_{n \rightarrow \infty} \frac{175}{1000}$

$$= \frac{175}{1000} = 0.175$$

(vi) P {outcome "6" for the experiment repeated is 1000 times} =  $\lim_{n \rightarrow \infty} \frac{190}{1000}$

$$= \frac{190}{1000} = 0.19$$

$$\text{Sum} = 0.179 + 0.15 + 0.157 + 0.149 + 0.175 + 0.19$$

$$= 1$$

Properties of Relative Frequency:

$0 \leq \frac{n(A)}{n} \leq 1$ , Relative frequency is between 0 & 1.

2) If  $\frac{n(A)}{n} = 0$ , then event "A" never occurs.

3) If  $\frac{n(A)}{n} = 1$ , then event "A" occurs in all "n"

4) If A & B are mutually exclusive events.   
 Any two events have no <sup>common</sup> <sup>element</sup> <sub>tail</sub>.

$$\text{Then } \frac{n(A \cup B)}{n} = \frac{n(A)}{n} + \frac{n(B)}{n}$$

Experiment :-

Experiment is a physical action & the process that is observed and the result is noted.

Ex:-

1. Tossing a coin
2. Rolling a single die.

Sample Space(s) :-

When a physical experiment is performed when the set of all the possible outcomes of a experiment is called a sample space.

\* It is denoted by "S"

\* Sample Space is universal set for the given experiment.

ex:-

(i) Tossing a coin has two outcomes head (H) and tail (T) so the sample space  $S = \{H, T\}$

(ii) Rolling a single die has 6 outcomes.

So, the "Sample" space  $S = \{1, 2, 3, 4, 5, 6\}$

## Discrete Sample Space:-

A sample space is said to be discrete and infinite if the set in the sample space have infinite number of elements.

ex:-  $N = \{1, 2, 3, 4, \dots, \infty\}$

## Continuous Sample Space:-

If the sample contains an infinite number of elements with continuous values with in a given range. Then it is called continuous sample space.

ex:-  $F = \{-0.5 \leq F \leq 8.5\}$  is an continuous sample space.

## Events:-

An event is defined as a set of possible outcomes of an experiment

\* Tossing a coin =  $\{H, T\}$

## Mutually Exclusive Events:-

If any two events in an experiment have no common outcomes then two events are said to be mutually exclusive events.

ex:- 1. Draw a King Card }  
2. Draw a Queen Card } are mutually exclusive events.

## Exhaustive events:-

All possible events in a sample space are called exhaustive events.



Ex: Tossing a coin  $\rightarrow$  Exhaustive events are  
-two Rolling a die  $\rightarrow$  Exhaustive events are  
Size.

Equally likely events:-

In a given experiment if one of the events is in sample it does not depend on the above another event. Then those events are called equally likely events.

ex: Tossing a coin getting heads and tails are equally likely events.

Independent Events:-

Let two events A & B in a sample space "S" if the probability of occurrence of events "A" is not effected by the probability of occurrence of event B. Then the two events are said to be two independent event.

Definition of probability:-

It is defined as Ratio of no. of favourable outcomes to the total number of possible outcomes.

$$P = \frac{\text{no. of favourable outcomes}}{\text{Total number of possible outcomes}}$$

Total number of possible outcomes.

Ex: When two dies are thrown find the probability sum getting 5.

Sol:

(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

Total possible outcomes = 36

Favourable outcomes for setting sum "5" is = 4

$$\text{Probability } p = \frac{4}{36} = \frac{1}{9}$$

Axioms of probability :-

i)  $P(A) \geq 0$ , the probability of A is greater than or equal to zero.

The probability of event "A" is always non negative real number.

ii)  $P(S) = 1$

The event "S" is an universal set and it is a sure event therefore the probability is always one.

$P(\emptyset) = 0$ , the null set  $\emptyset$  is an event and has no element it is known as impossible event and its probability is zero.

iii) If a Sample Space  $A_n$  has "N" element

where  $n = 1, 2, 3, 4, \dots, N$

If all the events are mutually exclusive events i.e.  $A_m \cap A_n = \emptyset$  for  $m \neq n$

$$P \left\{ \bigcup_{n=1}^N A_n \right\} = \sum_{n=1}^N P(A_n)$$

$$P(A_1) \cup P(A_2) \cup P(A_3) \dots \cup P(A_n) = P(A_1) + P(A_2) + P(A_3) \dots + P(A_n)$$

The theorem states that the probability of union of "N" number of mutually exclusive events is equal to the sum of probabilities of individual events.

Properties of probability:

1.  $0 \leq P(A) \leq 1$
2.  $P(\bar{A}) = 1 - P(A)$
3.  $P(S) = 1$
4.  $P(A) < P(B)$  if  $A \subset B$
5.  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

If  $A$  &  $B$  are mutually exclusive Events

Then  $P(A \cap B) = 0$

$$\therefore P(A \cup B) = P(A) + P(B)$$

Joint probability:

The Sample Space consists of two events  $A$  &  $B$  which are not mutually Exclusive then the

Probability of these events, occurs Simultaneously is called joint probability.

For two events  $A$  &  $B$  have same common element for the event  $A \cap B$  then for  $P(A \cap B)$  is called joint probability for the two events  $A$  &  $B$ .

$$P(A \cap B) = P(A) + P(B) - P(A \cup B)$$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

if  $P(A \cap B) \neq 0$  then

$$P(A \cup B) \leq P(A) + P(B)$$

conditional probability:-

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Conditional probabilities:  $P(A/B)$  The Conditional Probability  $P(A/B)$  is the probability of Event 'A' occurs on the condition that the probability of Event 'B' is already known.

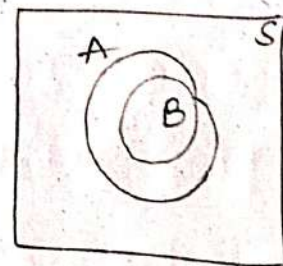
$$P(A/B) = \frac{P(A \cap B)}{P(B)}$$

Similarly the Conditional probability  $P(B/A)$  is the probability of Event 'B' occurs on the condition the probability of Event 'A' is already known.

$$P(B/A) = \frac{P(B \cap A)}{P(A)}$$

Properties of Conditional probability:

① For any two events A and B in a sample space 'S' if "B" is a subset of "A" then the probability of  $A/B = 1$



Proof:-

Given 'B' is a subset of 'A' as shown in figure the  $P(A \cap B) = P(B)$

we know that  $P(A/B) = \frac{P(A \cap B)}{P(B)}$

$$= \frac{P(B)}{P(B)} = 1$$

$$\therefore P(A|B) = 1$$

② If 'B' is a subset of 'A' then  $P(B|A) =$

$$\frac{P(B)}{P(A)}$$

Proof :-

B is a subset of 'A' as shown in the

figure then  $P(A \cap B) = P(B)$

$$\text{we know that } P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{P(B)}{P(A)}$$

③  $P(A|B) \geq 0$  that is the Conditional Probability is always a positive integer

Proof :- we know that  $P(A|B) = \frac{P(A \cap B)}{P(B)}$

If  $P(A \cap B)$  is greater than or equal to

$$\rightarrow \text{zero } P(A \cap B) \geq 0 \quad \& \quad P(B) \geq 0$$

$$\text{then } P(A|B) \geq 0$$

④ If two units A and B are in sample space 'S' then  $P(S|A) = P(S|B) = 1$  and

$$P(A|S) = P(A) \quad \text{and} \quad P(B|S) = P(B)$$

Proof: - Given 'A' and 'B' are in sample space 'S' then  $A \subseteq S$  and  $B \subseteq S$  then,

$$P(A \cap S) = P(A) ; P(B \cap S) = P(B)$$

$$P(S|A) = \frac{P(A \cap S)}{P(A)} = \frac{P(A)}{P(A)} = 1$$

$$P(S|B) = \frac{P(B \cap S)}{P(B)} = \frac{P(B)}{P(B)} = 1$$

$$P(A|S) = \frac{P(A \cap S)}{P(S)} = \frac{P(A)}{1} = P(A)$$

$$P(B|S) = \frac{P(B \cap S)}{P(S)} = \frac{P(B)}{1} = P(B)$$

⑤ If two events A and B are mutually exclusive events then the joint events  $A \cap C$  and  $B \cap C$  are also mutually exclusive events then the probability of  $P(A \cup B) =$

$$P(A|C) + P(B|C)$$

Proof: we know that the conditional probability

$$P(A \cup B | C) = \frac{P((A \cup B) \cap C)}{P(C)}$$

Now consider  $P((A \cup B) \cap C) = P(A \cap C) \cup P(B \cap C)$

By using  $x \cup y = x + y$

$$P(A \cup B | C) = \frac{P(A \cap C) \cup P(B \cap C)}{P(C)}$$

$$P(A \cup B | C) = \frac{P(A \cap C)}{P(C)} \cup \frac{P(B \cap C)}{P(C)}$$

Proof:- Given 'A' and 'B' are in sample space 'S' then  $A \subseteq S$  and  $B \subseteq S$ . Then,

$$P(A \cap S) = P(A) ; P(B \cap S) = P(B)$$

$$P(S|A) = \frac{P(A \cap S)}{P(A)} = \frac{P(A)}{P(A)} = 1$$

$$P(S|B) = \frac{P(B \cap S)}{P(B)} = \frac{P(B)}{P(B)} = 1$$

$$P(A|S) = \frac{P(A \cap S)}{P(S)} = \frac{P(A)}{1} = P(A)$$

$$P(B|S) = \frac{P(B \cap S)}{P(S)} = \frac{P(B)}{1} = P(B)$$

⑤ If two events A and B are mutually exclusive events then the joint events  $A \cap C$  and  $B \cap C$  are also mutually exclusive events then the probability of  $P(A \cup B) =$

$$P(A|C) + P(B|C)$$

Proof:- we know that the conditional probability

$$P(A \cup B | C) = \frac{P((A \cup B) \cap C)}{P(C)}$$

Now consider  $P((A \cup B) \cap C) = P(A \cap C) \cup P(B \cap C)$

By using  $(x+y) \cap z = x \cap z + y \cap z$

$$P(A \cup B | C) = \frac{P(A \cap C) \cup P(B \cap C)}{P(C)}$$

$$P((A \cup B) | C) = \frac{P(A \cap C) \cup P(B \cap C)}{P(C)}$$



$$P(A \cap B | C) = P(A|C) \cup P(B|C)$$

$$P[(A \cap B) | C] = P(A|C) + P(B|C)$$

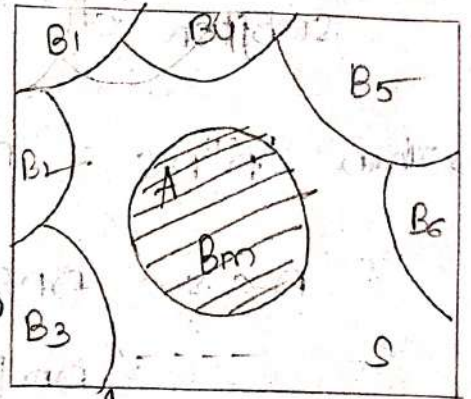
Hence proved.

Total probability theorem:

consider a sample space

"S" has "N" number of mutually exclusive events  $B_n$

$n=1, 2, 3, \dots, N$  such that  $B_m \cap B_n = \emptyset$



form  $n=1$  then the probability of any event

"A" defined all the sample space and can be expressed in terms of conditional

probability of event  $B_n$  this probability is called total probability of event 'A'.

$$\therefore P(A) = \sum_{n=1}^N P(A|B_n) P(B_n)$$

Proof: - The sample space "S" has "N" number

of mutually exclusive  $B_n$  where  $n = 1, 2, 3, \dots, N$ .

As shown in the figure

$\Rightarrow$  The events have the properties  $B_m \cap B_n = \emptyset$

$$= \{\emptyset\} \text{ form } n=1$$

$$\bigcup_{n=1}^{\infty} B_n = S \quad (B_1, B_2, B_3, \dots)$$

$$B_1 \cup B_2 \cup B_3 \cup \dots \cup B_n = S \quad (B_1, B_2, B_3, \dots)$$

$\Rightarrow$  Let "A" an even "A" be defined on the sample space "S" where the "A" is the subset of "S"

$$\Rightarrow \text{Then } A \cap S = A \subseteq S$$

$$A \subseteq S = A \cap S = A$$

$$= A \cap \left[ \bigcup_{n=1}^{\infty} B_n \right] = A$$

$$= A = \left[ \bigcup_{n=1}^{\infty} (A \cap B_n) \right]$$

Apply probability on b.s

$$P(A) = P \left[ \bigcup_{n=1}^{\infty} (A \cap B_n) \right]$$

$$P(A) = \bigcup_{n=1}^{\infty} P(A \cap B_n)$$

$$\frac{P(A \cap B)}{P(B)}$$

From the definition of joint probability

the probability of  $P(A \cap B_n) = P(A|B_n)P(B_n)$

$$P(A) = \bigcup_{n=1}^{\infty} P(A|B_n)P(B_n)$$

$$P(A) = \sum_{n=1}^{\infty} P(A|B_n)P(B_n)$$

$$P(A) = \sum_{n=1}^N P(A|B_n) P(B_n)$$

$$P(A) = \sum_{n=1}^N P(A|B_n) P(B_n)$$

### Bayes Theorem:

The Bayes's theorem states that if a sample space "S" has "N" mutually exclusive events  $B_n$  where  $n = 1, 2, 3, \dots, N$  such that

$B_m \cap B_n = \{\emptyset\}$  for  $m \neq n = 1, 2, 3, \dots, N$  and an event "A" is defined on the sample

space then the conditional probability of  $B_n$  and A is given by

$$P(B_n|A) = \frac{P(A|B_n) P(B_n)}{P(A|B_1) P(B_1) + P(A|B_2) P(B_2) + \dots + P(A|B_N) P(B_N)}$$

$$P(A|B_1) P(B_1) + P(A|B_2) P(B_2) + \dots + P(A|B_N) P(B_N)$$

### Proof:

We know that the conditional probability

$$P(B_n|A) = \frac{P(A \cap B_n)}{P(A)} \text{ if } P(A) \neq 0 \text{ --- (1)}$$

$$\text{Alternatively } P(A|B_n) = \frac{P(A \cap B_n)}{P(B_n)} \text{ if } P(B_n) \neq 0 \text{ --- (2)}$$

From eq. (2)

$$P(A \cap B_n) = P(A|B_n) P(B_n) \text{ --- (3)}$$

One form of Bayes's theorem is obtained by

substituting eq. (3) in eq. (1)

$$P(B_n|A) = \frac{P(A|B_n)P(B_n)}{P(A)} \quad \text{--- (4)}$$

we know that the total probability theorem,  $P(A) = \sum_{n=1}^N P(A|B_n)P(B_n)$  --- (5)

Sub eq (5) in eq (4)

$$P(B_n|A) = \frac{P(A|B_n)P(B_n)}{\sum_{n=1}^N P(A|B_n)P(B_n)}$$

$$P(B_n|A) = \frac{P(A|B_n)P(B_n)}{P(A|B_1)P(B_1) + P(A|B_2)P(B_2) + \dots + P(A|B_N)P(B_N)}$$

### Problems:-

1. In a box there are 100 resistors having resistance and tolerance values as shown in table. Let a resistor be selected from the box and assume that each resistor has the same likelihood of chosen.

Resistance	Tolerance		Total
	5%	10%	
22-Ω	10	14	24
47-Ω	28	16	44
100-Ω	20	9	29

Event A: Draw a  $47\Omega$  resistor

Event B: Draw a resistor with 5% tolerance

Event C: Draw a  $100\Omega$  resistor

Find (i) Individual probability  $\rightarrow P(A), P(B), P(C)$

(ii) Joint probability  $\rightarrow P(A \cap B), P(B \cap C), P(A \cap C)$

(iii) Conditional probability.  $\rightarrow P(A|B), P(B|A), P(C|A), P(A|C), P(C|B), P(B|C)$

(i) Individual probabilities

$\Rightarrow$  Given Event "A" is draw a  $47\Omega$  resistor

then the probability of A  $P(A) = \frac{44}{100} = 0.44$

$\Rightarrow$  The Event "B" draw a resistor 5%

tolerance then  $P(B) = \frac{62}{100} = 0.62$

$\Rightarrow$  The Event "C" draw a  $100\Omega$  resistor

the  $P(C) = \frac{32}{100} = 0.32$

(ii) Joint probability

$\Rightarrow$  The joint Event "AB" is draw a  $47\Omega$  resistor with 5% tolerance. There are 28 such resistors.

$$\therefore P(A \cap B) = \frac{28}{100} = 0.28$$

$\Rightarrow$  The joint Event "AC" is draw a  $47\Omega$  resistor

with 100-2 resistors.

$$\therefore P(AnC) = 0$$

Because there are no such resistors.

$$P(BnC) = \frac{24}{100} = 0.24.$$

(iii) Conditional probability:-

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{0.28}{0.62}$$

$$P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{0.28}{0.44}$$

$$P(C|A) = \frac{P(A \cap C)}{P(A)} = \frac{0}{0.44} = 0.$$

$$P(A|C) = \frac{P(A \cap C)}{P(C)} = 0$$

$$P(B|C) = \frac{P(B \cap C)}{P(C)} = \frac{0.24}{0.32} = 0.75$$

$$P(C|B) = \frac{P(B \cap C)}{P(B)} = \frac{0.24}{0.62} = 0.38$$

Note:-

When two events A and B are independent events then  $P(A \cap B) = P(A)P(B)$

$$P(A \cap B \cap C) = P(A)P(B)P(C).$$

Problems

From a pack of 52 cards define event

"A" drawing a king card event "A" and

Event "B" drawing a jack & Queen cards and

Event "C" drawing a heart card. then  
find which are independent and dependent  
Events.

Sol<sup>o</sup> A pack of 52 cards.

Then the probability of drawing a king  
card

$$\text{i.e. } P(A) = \frac{4}{52} \quad (\text{since there are 4 king cards}).$$

Then the probability of drawing a jack &  
Queen card

$$\text{i.e. } P(B) = \frac{8}{52} \quad (\text{since there are 4 jack \& 4 Queen cards})$$

Then the probability of drawing a heart  
card i.e.  $P(C) = \frac{13}{52}$  (since there are 13 heart  
cards)

$$P(A \cap B) = 0$$

Since A & B are disjoint events.

$$P(A) \cdot P(B) = \frac{4}{52} \cdot \frac{8}{52} = \frac{32}{2704} = 0.0118$$

$$\therefore P(A \cap B) \neq P(A)P(B)$$

Then two events A & B dependent events.

$P(B \cap C) = \frac{2}{52}$  (since there are one jack and one  
Queen in heart card).

$$P(B) \cdot P(C) = \frac{8}{52} \cdot \frac{13}{52} = \frac{104}{2704} = \frac{1}{26}$$

$$P(A \cap C) = P(A) \cdot P(C)$$

$\therefore$  Then the two events A & B are independent events.

$$P(A \cap C) = \frac{1}{58}$$

$$P(A) \cdot P(C) = \frac{4}{58} \cdot \frac{13}{58} = \frac{52}{3364} = \frac{1}{64.7}$$

$$\therefore P(A \cap C) = P(A) \cdot P(C)$$

$\therefore$  Then the two events A & B are independent event.

In a factory there are four machines, the machines produce 10%, 20%, 30% and 40% of an items respectively. The defective items to be produced by each machine are 5%, 4%, 3% and 2% respectively.

Now an item is selected which is too be defective and what is the probability of defective item being from the second machine.

In a factory given 4 machines to produce an items.

Let  $B_1, B_2, B_3$  and  $B_4$  are the events of

to producing an items by the 4 machines.

$$\text{Then } P(B_1) = \frac{10}{100} = 0.1$$



$$P(B_3) = \frac{30}{100} = 0.3$$

$$P(B_4) = \frac{40}{100} = 0.4$$

Let "D" be the event to producing a defective item.

Given the probability of two produce items by each machine are  $P(D|B_1) = \frac{5}{100} = 0.05$

$$P(D|B_2) = \frac{4}{100} = 0.04$$

$$P(D|B_3) = \frac{3}{100} = 0.03$$

$$P(D|B_4) = \frac{2}{100} = 0.02$$

Using total probability theorem defective items to be produced by all the machines

$$\text{are } P(D) = \sum_{N=1}^4 P(D|B_N) P(B_N)$$

$$P(D) = P(D|B_1)P(B_1) + P(D|B_2)P(B_2) + P(D|B_3)$$

$$P(B_3) + P(D|B_4)P(B_4)$$

$$P(D) = 0.05(0.1) + 0.04(0.2) + 0.03(0.3) + 0.02(0.4) \\ = 0.03$$

The probability of getting defective items to be produced by second machine,

$$P(B_2|D) = \frac{P(D|B_2)P(B_2)}{P(D)} \quad [\because \text{Baye's Theorem}]$$

$$= \frac{0.04(0.2)}{0.07} = 0.286$$

Problems:-

① When two dice are thrown, determine the probabilities for the following three events.

(i)  $A = \text{Sum} = 7$  (ii)  $B = \{8 < \text{Sum} \leq 11\}$  (iii)  $C = \{10 < \text{Sum}\}$   
and also determine (iv)  $P(B \cap C)$  (v)  $P(B \cup C)$

Sol: When two dice are thrown

total outcomes are  $6^2 = 36$

(i) For the event  $A = \{\text{Sum} = 7\}$ , the favourable outcomes are  $\{(6, 1), (1, 6), (5, 2), (2, 5), (4, 3), (3, 4)\}$

$$\therefore P(A) = P(\text{Sum} = 7) = \frac{6}{36} = \frac{1}{6} = 0.167$$

(ii) For the event  $B = \{8 < \text{Sum} \leq 11\} = \{\text{Sum} = 9, 10 \text{ or } 11\}$

The favourable outcomes are  $\{(3, 6), (6, 3), (4, 5), (5, 4), (6, 4), (4, 6), (5, 5), (5, 6), (6, 5)\}$

$$P(B) = P\{8 < \text{Sum} \leq 11\} = \frac{9}{36} = \frac{1}{4} = 0.25$$

(iii) For the event  $C = \{10 < \text{Sum}\} = \{\text{Sum} = 11 \text{ or } 12\}$

The favourable outcomes are  $\{(5, 6), (6, 5), (6, 6)\}$

$$P(C) = P\{10 < \text{Sum}\} = \frac{3}{36} = \frac{1}{12} = 0.083$$

(iv) For the event  $B \cap C$ , the outcomes of common to both the events B and C are  $\{(5, 6), (6, 5)\}$

$$\therefore P(B \cap C) = \frac{2}{36} = \frac{1}{18} = 0.055$$

(v)  $P(BUC)$

For the event  $BUC$ , the outcomes of event

$$\text{are } P(BUC) = P(B) + P(U) - P(B \cap U)$$

$$= \frac{9}{36} + \frac{3}{36} - \frac{2}{36} = \frac{10}{36} = 0.278.$$

Q2 Two boxes are selected randomly. The first box contains 2 white balls and 3 black balls. The second box contains 3 white and 4 black balls. What is probability of drawing a white ball?

Sol: Assume that two boxes have equal probability of being selected.

The probability of selecting the first box is  $B_1$ .

The probability of selecting the 2nd box is  $B_2$ .

$$P(B_1) = P(B_2) = \frac{1}{2} = 0.5$$

Given the favourable event is "selecting white ball".

The probability of selecting white ball from

box  $B_1$  is  $P(W|B_1) = \frac{2}{5} = 0.4$

The probability of selecting white ball from

box  $B_2$  is  $P(W|B_2) = \frac{3}{7} = 0.4285$ .

Using the total probability theorem.

$$P(W) = P(W|B_1)P(B_1) + P(W|B_2)P(B_2)$$

$$= \frac{1}{2} \left( \frac{2}{5} + \frac{3}{7} \right) = \frac{29}{70} = 0.4143.$$

③ A shipment of components contains of three identical boxes. One box contains 2000 components of which 25% are defective. The second has 5000 components of which 20% are defective and the third box contains 2000 components of which 60% are defective. A box is selected randomly and a component is removed at random from the box. What is the probability that this component is defective? What is the probability that it came from the second box.

Given 3 identical boxes. Let the events  $B_1$ ,  $B_2$  and  $B_3$  are the selections of boxes. Assume that the selection is equally likely  $P\{B_i\}$ .

$$P\{B_1\} = P\{B_2\} = P\{B_3\} = \frac{1}{3} = 0.33$$

Let the event "D" be "selecting a defective component".

Given 25% of the components are defective in

$B_1$

$$\therefore P(D|B_1) = \frac{25}{100} = 0.25$$

20% of the components are defective in  $B_2$ ,

$$P(D|B_2) = \frac{20}{100} = 0.2$$

60% components are defective from 2000 in  $B_3$

$$P(D|B_3) = \frac{600}{1000} = \frac{3}{5} = 0.6$$

using total probability theorem, the probability of getting ~~total probability~~ a defective component is

$$P(D) = P(D|B_1)P(B_1) + P(D|B_2)P(B_2) + P(D|B_3)P(B_3) \\ = (0.25 + 0.2 + 0.3) \frac{1}{3} = \frac{0.75}{3} = 0.25$$

The probability that the defective component comes from the second box is

$$P(B_2|D) = \frac{P(D|B_2)P(B_2)}{P(D)} \text{ (using Bayes' theorem)} \\ = \frac{0.2 + \frac{1}{3}}{0.25} = \frac{20}{25} = 0.267.$$

④ A letter is known as to have come from either "TATANAGAR" or "CALCUTTA" on the envelope... Just two consecutive letters, "TA" are visible. Find the probability that the letter has come from "CALCUTTA".

Sol: let the event "A" be letter from TATANAGAR  
let the event "B" be letter from CALCUTTA.

assume that both events are equally likely  $P(A) = P(B) = \frac{1}{2}$ .

let the event "C" be the two consecutive letters "TA".

The probability of the ~~two~~ consecutive letters "TA" being selected from TATANAGAR is

TA TA N A G AR

$$P(C|A) = \frac{2}{7}$$

The probability of the letters "TA" being selected from CALCUTTA is CALCUTTA

$$P(C|B) = 1/7$$

Using Bayes' theorem, the probability of the letter coming from "CALCUTTA" when the event "c" selected is

$$P(B|C) = \frac{P(C|B)P(B)}{P(C|A)P(A) + P(C|B)P(B)}$$

$$P(B|C) = \frac{\frac{1}{7} \times \frac{1}{2}}{\left(\frac{2}{7} \times \frac{1}{2}\right) + \left(\frac{1}{7} \times \frac{1}{2}\right)}$$

$$= \frac{\frac{1}{7}}{\frac{2}{7} + \frac{1}{7}}$$

$$= 0.33$$

THE RANDOM VARIABLE

The primary goal of this subject are to introduce the principles of random signals and to provide tools one can deal with systems involving such signals.

- A random signal is a time waveform that can be characterized only in some probabilistic manner. It can be either a desired or undesired waveform.
- The reader has heard background hiss while listening to an ordinary broadcast radio receiver. The waveform causing the hiss, when observed on an oscilloscope, would appear as a randomly fluctuating voltage with time, it is undesirable.
- Since it interferes with our ability to hear the radio program, and is called noise.
- In a radio astronomer's receiver, noise interferes with the desired signal from outer space.
- In a television system, noise shows up in the form of picture interference often called "snow".
- In a sonar system, randomly generated sea sounds give rise to a noise that interferes with the desired echoes.
- The bits in a computer bit stream appear to fluctuate randomly with time b/w the zero and one states thereby creating randomness.
- The o/p voltage of a wind power generator would be random because wind speed fluctuates randomly.
- The voltage from a sonar detector varies randomly due to the randomness of cloud and weather conditions.

→ A set is a collection of objects. The objects are called elements of the set. A set is usually denoted by capital letter while an element is represented by a lower case letter.  $a \in A$ .

A set is specified by the content of two braces:  $\{ \}$ .

→ A set is said to be countable, if the elements can be put in one-to-one correspondence with natural numbers:  $A = \{1, 2, 3, \dots\}$

→ If a set is not countable, it is called uncountable  
 $C = \{0.5 < C \leq 8.5\}$

→ A set is said to be empty, if it has no elements. The empty set is given the symbol  $\emptyset$  and is often called null set.

→ A finite set is one that is either empty or has elements that can be counted. (B) It has a finite no. of elements.

$D = \{0, 0\}$ ,  $E = \{2, 4, 6, 8, 10, 12, 14\}$

→ If a set is not countable, ~~it~~ finite, it is called infinite. An infinite set having countable elements is called countable infinite  $B = \{1, 2, 3, \dots\}$

An infinite set having uncountable elements is called uncountable infinite  $F = \{-5.6 < F \leq 12.0\}$ .

→ Experiment is any physical action or process that is observed and the result is noted. EX- Tossing a coin, Firing a missile

→ A single performance of the experiment is called a trial for which there is an outcome.

→ The set of all possible outcomes in any given experiment is called the sample space. and it is given by the symbol "S".

→ An event is defined as a subset of sample space

EX:- when die is thrown the sample space  $S = \{1, 2, 3, 4, 5, 6\}$

$A = \{1, 3, 5\}$  is the event of getting an odd number

$B = \{2, 4, 6\}$  " " " even "

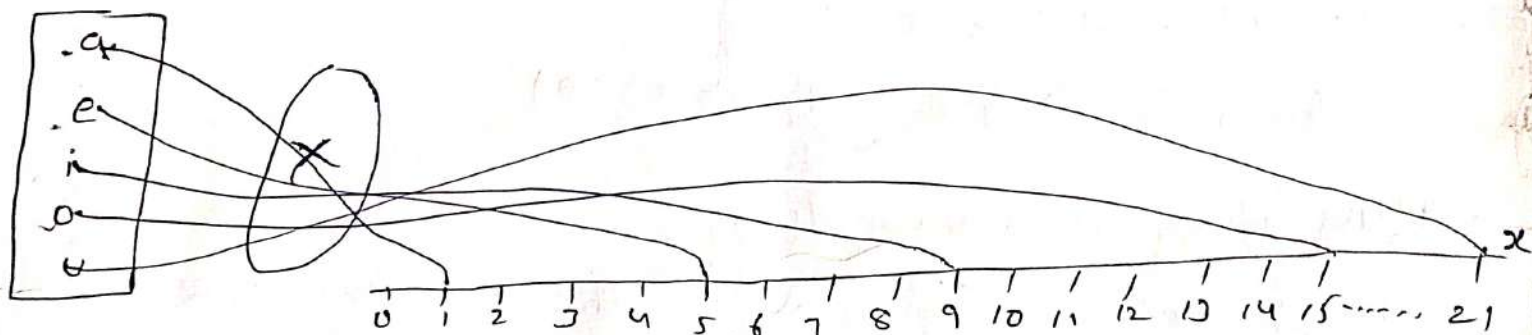
both  $A \subseteq S$ ,  $B \subseteq S$



(2)

we define a real random variable as a real function of the elements of a sample space "S". we shall represent a random variable by a capital letter (such as  $W, X, Y$ ) and any particular value of the random variable by a lowercase letter (such as  $w, x, y$ ). Thus given an experiment defined by a sample space "S" with elements  $s$ , we assign to every  $s$  a real number, according to some rule and call  $X(s)$  a random variable.

Example: consider a sample space "S" consisting of alphabets  $a, e, i, o, u$ . we take a random variable 'X' in such a way that each outcome corresponds to a real number.



Here the letter 'a' is the first letter in English letters and hence the real number to be assigned is one. similarly 'e' is assigned to '5', and 'i' is assigned to 9 and 'o' is assigned to the mark 15 and 'u' is assigned to the number 21. Thus the random variable 'X' can be considered to a function that maps all the elements of sample space into points on the real line.

conditions for a function to be a random variable :-

The following are the necessary conditions to be satisfied by a random variable

condition I :- It is not necessary that the sample space points map uniquely i.e. more than outcome may map to a

single real value

eg:  $X(A_1) = 2$  and  $X(A_2) = 2$

But the random variable must not be multivalued i.e. every point on the sample space must correspond to only one real value of R.V

Condition 2: The set  $\{X \leq x\}$  shall be an event for any real number 'x' i.e. the prob of this event denoted by  $P\{X \leq x\}$  is equal to the sum of the probabilities of all the elementary events corresponding to  $\{X \leq x\}$

eg:  $P\{X \leq 4\} = P\{X=4\} + P\{X=3\} + P\{X=2\} + P\{X=1\} + P\{X=0\}$ .

→ The condition we require that the probabilities of the events  $\{X = -\infty\}$  and  $\{X = \infty\}$  be 0

i.e.  $P\{X = -\infty\} = 0$  &  $P\{X = \infty\} = 0$

Types of random variables :-

Basically random variables are classified into two types  
1. discrete random variable 2. continuous random variable.  
and also another random variable called as mixed random

Discrete random variable :-

In any finite interval, if a random variable assumes only finite no. of distinct values then random variable is called discrete random variable.

eg: no. of telephone calls arriving at an office in a finite no. of interval of time.

Continuous random variable :-

If a random variable assumes infinite no. of uncountable values then it is called continuous random variable.

eg - falling of a rain drops

## Mixed random variable!.

(12)

A mixed random variable is one for which some of its values are discrete and some are continuous.

Problem:- An experiment consists of a rolling a die and flipping a coin. The random variable  $x$  chosen such that

(i) a coin head (H) outcome, corresponds to positive values of  $x$  that are equal to the numbers that show up on the die.

(ii) A coin tail (T) outcome corresponds to negative values of  $x$  that are equal in magnitude to twice the numbers that shows on the die. Map the elements of random variable " $x$ " into points on the real line and explain.

Sol: Given that a coin head outcome corresponds to the +ve values of " $x$ " which are equal to the numbers that show up on the side.

We know that the outcomes of die  $S = \{1, 2, 3, 4, 5, 6\}$ .

The values  $x$  for head outcomes are

$$(H,1) = 1; (H,2) = 2; (H,3) = 3; (H,4) = 4; (H,5) = 5; (H,6) = 6;$$

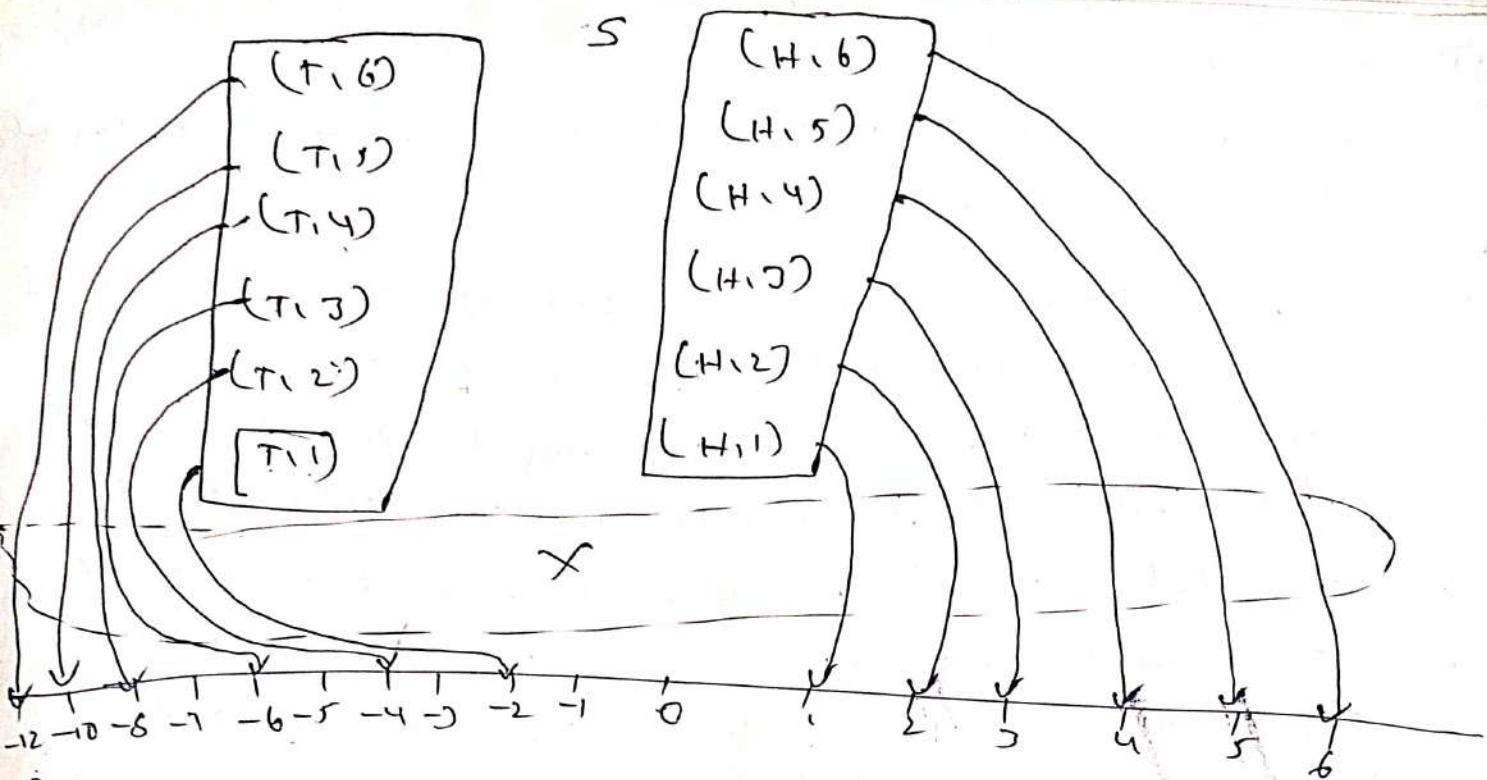
Given the coin tail (T) outcome corresponds to the -ve values of  $x$  which are equal in magnitude to twice the numbers that shown up on a die.

The values of  $x$  for tail outcome is

$$(T,1) = -2, (T,2) = -4, (T,3) = -6, (T,4) = -8, (T,5) = -10,$$

$$(T,6) = -12$$

Explanation:- Hence for a head (H) on the coin the values of random variable  $x$  lies on the +ve side of the real line. For tail on the coin the values of random variable " $x$ " lies on the -ve side of real line.



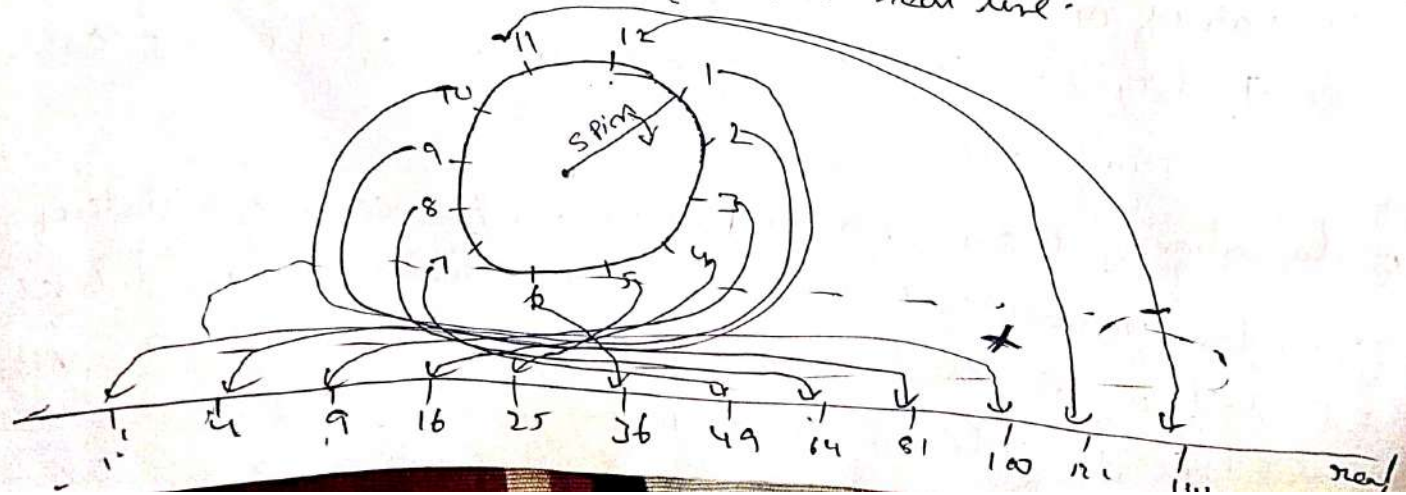
(2)

In experiment where the pointer on wheel of chance spins. The possible outcomes are the numbers from 0 to 12 marked on the wheel. The sample space consists of the numbers on the set  $0 \leq S \leq 12$ . and if the random variable defined as  $X = X(S) = S^2$ . map the elements of random variable on real line & explain.

Given that sample space  $S = \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12\}$ .  
 Given  $X(S) = S^2$

S	1	2	3	4	5	6	7	8	9	10	11	12
$X = S^2$	1	4	9	16	25	36	49	64	81	100	121	144

Here we can mapping the points of wheel i.e sample space with values of random variable "X" on a real line.



(4)

Probability distribution function [cumulative distribution function (CDF)]

Probability distribution function is defined as

$$F_X(x) = P\{X \leq x\}, \quad -\infty \leq x \leq \infty$$

The probability  $P\{X \leq x\}$  is the probability of the event  $\{X \leq x\}$ .

Here 'X' is the function of random variable and x is the values of random variable. i.e. it is a function of x.

EX:- consider a rolling die  $\xi_1$ , X be the random variable which indicates the faces of die with numbers.

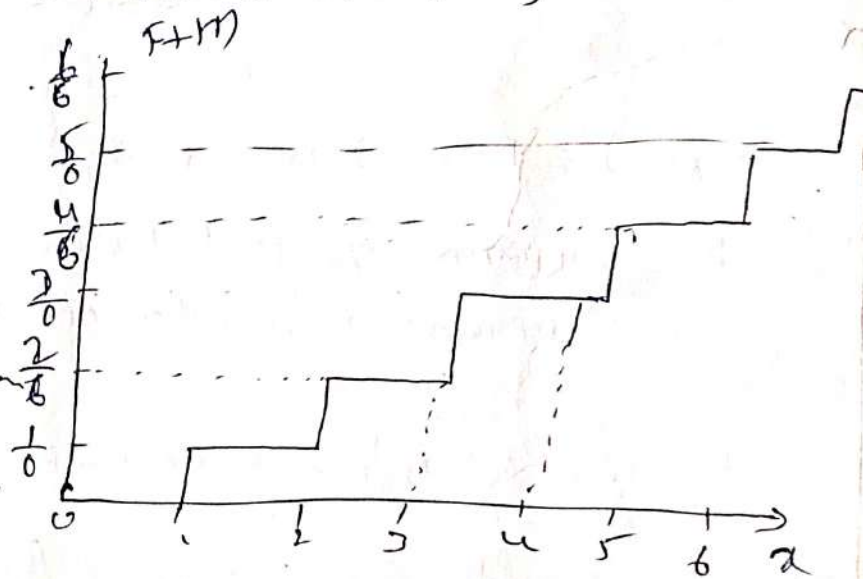
$x = x_i$	1	2	3	4	5	6
$P(x_i)$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$

From the definition of distribution function  $F_X(x) = P\{X \leq x\}$ .

Hence  $F_X(1) = P\{X \leq 1\} = P(X=1) = \frac{1}{6}$

$$F_X(2) = P\{X \leq 2\} = P(X=1) + P(X=2) = \frac{1}{6} + \frac{1}{6} = \frac{2}{6}$$

The probability distribution function is in the form of stairs case as shown in fig.



NOTE:- The expression for distribution function for a discrete random variable can be written as

$$F_X(x) = \sum_{i=1}^N P(x_i) u(x - x_i)$$

Here  $u(x)$  denotes the unit step function and is given by

$$u(x) = \begin{cases} 1 & x \geq 0 \\ 0 & x < 0 \end{cases}$$

## Properties of CDF:-

1.  $F_X(-\infty) = 0$  and  $F_X(\infty) = 1$

Proof:- From the definition  $F_X(x) = P\{X \leq x\}$

$$\text{Hence } F_X(-\infty) = P\{X \leq -\infty\}$$

There are no real values exists less than  $-\infty$  hence  $F_X(-\infty) = 0$

$$\text{Similarly } F_X(\infty) = P\{X \leq \infty\}$$

$$= P\{S\} = 1$$

$$\therefore \underline{F_X(\infty)} = 1 \text{ \& } F_X(-\infty) = 0$$

2.  $0 \leq F_X(x) \leq 1$

Proof:- We know that  $F_X(-\infty) = 0$  and  $F_X(\infty) = 1$

hence  $F_X(x)$  lies b/w 0 and 1

$$\therefore 0 \leq F_X(x) \leq 1$$

3.  $F_X(x_1) \leq F_X(x_2)$  if  $x_1 < x_2$

This property says that the distribution function  $F_X(x)$  is non-decreasing function of "x".

4.  $P\{x_1 < X \leq x_2\} = F_X(x_2) - F_X(x_1)$

Proof:- Consider  $\{X \leq x_1\}$  and  $\{x_1 < X \leq x_2\}$  are mutually exclusive events.  $\therefore \{X \leq x_2\} = \{X \leq x_1\} + \{x_1 < X \leq x_2\}$

$$P\{X \leq x_2\} = P\{X \leq x_1\} + P\{x_1 < X \leq x_2\}$$

$$F_X(x_2) = F_X(x_1) + P\{x_1 < X \leq x_2\}$$

$$\therefore P\{x_1 < X \leq x_2\} = F_X(x_2) - F_X(x_1)$$

$$\rightarrow F_x(x^+) = F_x(x)$$

5

This property says that distribution function is continuous from right.

$$\rightarrow P\{x > x\} = 1 - F_x(x)$$

Proof: consider  $\{x \leq x\}$  and  $\{x > x\}$  are mutually exclusive events

$$\therefore \{x > x\} + \{x \leq x\} = S$$

$$P\{x > x\} + P\{x \leq x\} = P(S)$$

$$P\{x > x\} = 1 - F_x(x)$$

Problem:

A random variable 'x' has the following probab function

value of x :	0	1	2	3	4	5	6
P(x) :	k	3k	5k	7k	9k	11k	13k

Find (i) the value of 'k' (ii)  $P(x < 4)$ ,  $P(x \geq 5)$  and  $P(3 \leq x \leq 6)$

- (iii) what is the smallest value of x for which  $P\{x \leq x\} = \frac{1}{2}$   
 (iv) what is the min value of k so that prob of  $P\{x \leq 2\} > 0$   
 (v) Find the distribution function,  $F_x(x)$ .

Sol: (i) we know that the total probability is always unity

$$\sum_{i=0}^6 P(x_i) = 1 \Rightarrow k + 3k + 5k + 7k + 9k + 11k + 13k = 1$$

$$k = \frac{1}{49}$$

$$(ii) P\{x < 4\} = P(x=0) + P(x=1) + P(x=2) + P(x=3)$$

$$= k + 3k + 5k + 7k = 16k = \frac{16}{49}$$

$$P\{x \geq 5\} = P(x=5) + P(x=6) = 11k + 13k = 24k = \frac{24}{49}$$

$$P\{3 \leq x \leq 6\} = P(x=4) + P(x=5) + P(x=6) = 9k + 11k + 13k = \frac{32}{49}$$

$$(iii) \text{ consider } P\{x \leq 0\} = P(x=0) = k = \frac{1}{49}$$

$$P\{x \leq 1\} = P(x=0) + P(x=1) = \frac{1}{49} + \frac{3}{49} = \frac{4}{49}$$

$$P\{x \leq 2\} = P(x=0) + P(x=1) + P(x=2) = \frac{9}{49} = 0.18$$

$$P\{x \leq 3\} = P(x=0) + P(x=1) + P(x=2) + P(x=3) = \frac{16}{49} = 0.32$$

$$P\{x \leq 4\} = P(x=0) + P(x=1) + P(x=2) + P(x=3) + P(x=4) = \frac{25}{49} = 0.51 > 0.5$$

Hence smallest value of  $x_i$  is 4 for  $P\{x \leq x\} > \frac{1}{2}$ .

(iv)  $P\{x \leq 2\} \geq 0.3$

$$P(x=0) + P(x=1) + P(x=2) > 0.3$$

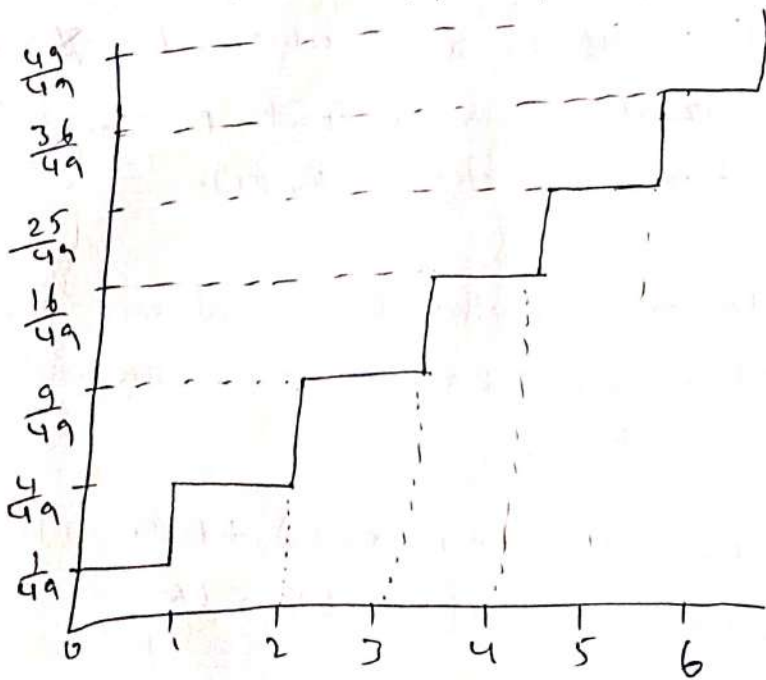
$$k + 3k + 5k > 0.3$$

$$9k > 0.3$$

$$k > \frac{0.3}{9} = \frac{1}{30}$$

The min value of  $k$  is  $\frac{1}{30}$

$x_i$	0	1	2	3	4	5	6
$P(x_i)$	$\frac{1}{49}$	$\frac{3}{49}$	$\frac{5}{49}$	$\frac{7}{49}$	$\frac{9}{49}$	$\frac{11}{49}$	$\frac{13}{49}$
$F+(x)$	$\frac{1}{49}$	$\frac{4}{49}$	$\frac{9}{49}$	$\frac{16}{49}$	$\frac{25}{49}$	$\frac{36}{49}$	$\frac{49}{49}$



NOTE: The expression of distribution function for discrete random variable is given by  $F+(x) = \sum_{i=0}^{\infty} P(x_i) u(x-x_i)$ .



→ The random variable 'x' has the discrete variable on the set  $\{-1, -0.5, 0.7, 1.5, 3\}$  the corresponding probabilities are assumed to be  $\{0.1, 0.2, 0.1, 0.4, 0.2\}$ . Plot its distribution function and state its discrete or continuity.

Sol

$x_i$	$-1$	$-0.5$	$0.7$	$1.5$	$3$
$P(x_i)$	$0.1$	$0.2$	$0.1$	$0.4$	$0.2$

we know that the probability distribution function can be defined as

$$F(x) = P\{x \leq x\}$$

$$F(-1) = P\{x \leq -1\} = 0.1,$$

$$F(-0.5) = P\{x \leq -0.5\} = 0.1 + 0.2 = 0.3$$

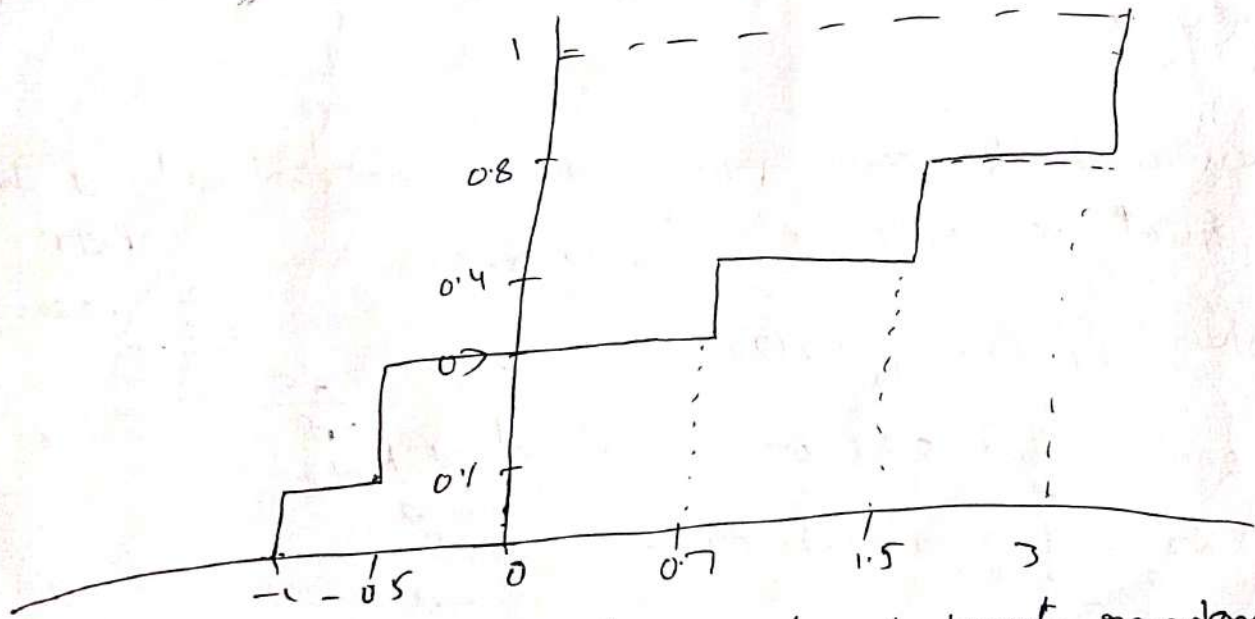
$$F(0.7) = P\{x \leq 0.7\} = P(x=-1) + P(x=-0.5) + P(x=0.7)$$

$$= 0.1 + 0.2 + 0.1 = 0.4$$

$$F(1.5) = 0.4 + 0.4 = 0.8$$

$$F(3) = 0.8 + 0.2 = 1$$

$x_i$	$-1$	$-0.5$	$0.7$	$1.5$	$3$
$P(x_i)$	$0.1$	$0.2$	$0.1$	$0.4$	$0.2$
$F(x)$	$0.1$	$0.3$	$0.4$	$0.8$	$1$



The expression for distribution function of discrete random variable

$$F(x) = \sum P(x_i) u(x-x_i)$$

$$F(x) = 0.1 u(x+1) + 0.2 u(x+0.5) + 0.1 u(x+0.7) + 0.4 u(x-1.5) + 0.2 u(x-3)$$

## Probability density function (P.D.F) ( $f_x(x)$ ) :-

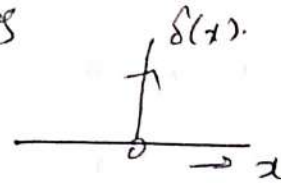
The probability density function can be defined as the derivative of distribution function  $F_x(x)$ . and which is denoted by  $f_x(x)$  and is given by  $f_x(x) = \frac{d}{dx} F_x(x)$ .

we know that distribution function is in the form of stair case form with step function, unit impulse function can be used to indicate the derivative of unit stair function i.e.

$$\delta(x) = \frac{d}{dx} u(x)$$

Here  $\delta(x)$  is unit impulse function and  $u(x)$  is unit step function, unit impulse function can be defined as

$$\delta(x) = \begin{cases} 1 & x=0 \\ 0 & \text{elsewhere} \end{cases}$$



## Properties of probability density function :-

Property 1 :-  $f_x(x) \geq 0$  for all "x"

This property states that the probability density function  $f_x(x)$  is non-negative

→ Property 2

The area under the ~~curve~~ P.D.F curve over an interval  $-\infty$  to  $\infty$  is equal to unity i.e.  $\int_{-\infty}^{\infty} f_x(x) dx = 1$

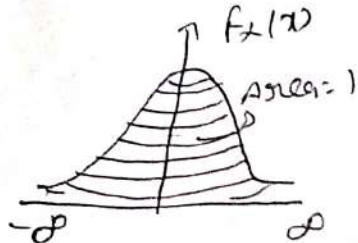
Proof: consider  $f_x(x) = \frac{d}{dx} F_x(x)$

Take integral on both sides over an interval  $-\infty$  to  $\infty$

$$\int_{-\infty}^{\infty} f_x(x) dx = \int_{-\infty}^{\infty} \frac{d}{dx} F_x(x) dx = F_x(x) \Big|_{-\infty}^{\infty}$$

$$\int_{-\infty}^{\infty} f_x(x) dx = F_x(\infty) - F_x(-\infty) = 1 - 0 = 1$$

$$\therefore \int_{-\infty}^{\infty} f_x(x) dx = 1$$



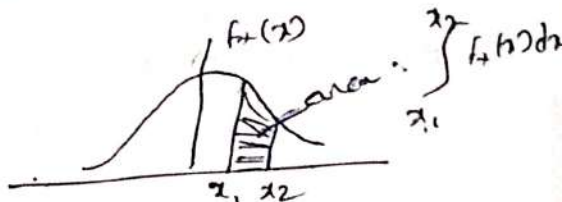
NOTE: This property is useful for check the validity of density of well as distribution function

Property 3 :-

$$F_x(x) = \int_{-\infty}^x f_x(x) \cdot dx = P\{x \leq x\}$$

This property states that distribution function equals to the integral of density function over an interval  $-\infty$  to  $x$ . (or)

The distribution function equals to the area under the P.D.F curve over an interval  $-\infty$  to  $x$ .



→ (4)  $P\{x_1 < x \leq x_2\} = \int_{x_1}^{x_2} f_x(x) \cdot dx$

NOTE :- If 'x' is a discrete random variable then

$$F_x(x) = \sum_{i=1}^N P(x_i) u(x-x_i)$$

$$f_x(x) = \sum_{i=1}^N P(x_i) \delta(x-x_i)$$

→ If 'x' is a continuous random variable then

$$F_x(x) = \int_{-\infty}^x f_x(x) dx$$

$$f_x(x) = \frac{d}{dx} F_x(x)$$

Problem :- Consider an experiment of tossing four fair coins. The random variable 'x' is associated with the no. of tails showing. compute and sketch the cumulative distribution function.

Sol :- Give 4 fair coins are tossed then possible outcomes are  $2^4 = 16$

- S = { HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH, HHHH }

Given x is a random variable associated with no. of tails :-

$$x = \{0, 1, 1, 1, 2, 1, 2, 2, 3, 1, 2, 2, 3, 2, 3, 3, 4\}$$

$$x = 0 \quad 1 \quad 2 \quad 3 \quad 4$$

$$P(x=x_i) = \frac{1}{16} \quad \frac{4}{16} \quad \frac{6}{16} \quad \frac{4}{16} \quad \frac{1}{16}$$

$$F_x(x) = \frac{1}{16} \quad \frac{5}{16} \quad \frac{11}{16} \quad \frac{15}{16} \quad \frac{16}{16}$$

we know that  $F_X(x) = P\{x \leq x\}$

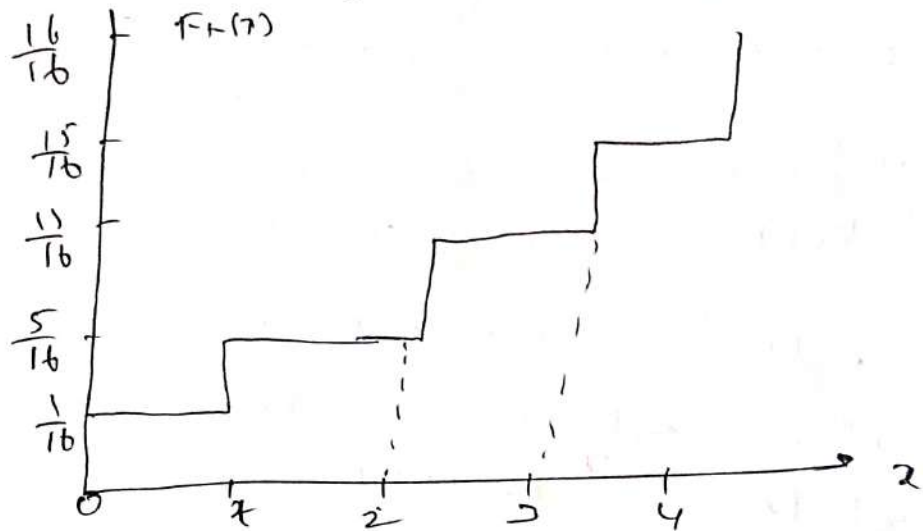
$$F_X(0) = P\{x \leq 0\} = P\{x=0\} = \frac{1}{16}$$

$$F_X(1) = P\{x \leq 1\} = P\{x=0\} + P\{x=1\} = \frac{5}{16}$$

$$F_X(2) = P\{x \leq 2\} = P\{x=0\} + P\{x=1\} + P\{x=2\} = \frac{11}{16}$$

$$F_X(3) = P\{x \leq 3\} = P\{x=0\} + P\{x=1\} + P\{x=2\} + P\{x=3\} = \frac{15}{16}$$

$$F_X(4) = P\{x \leq 4\} = P\{x=0\} + P\{x=1\} + P\{x=2\} + P\{x=3\} + P\{x=4\} = \frac{16}{16}$$



NOTE: From the above plot the CDF expression for discrete random variable can be written as

$$F_X(x) = \sum_{i=0}^n P(x_i) u(x - x_i)$$

$$F_X(x) = \frac{1}{16} u(x) + \frac{4}{16} u(x-1) + \frac{6}{16} u(x-2) + \frac{4}{16} u(x-3) + \frac{1}{16} u(x-4)$$

⇒ A random variable "X" has the probabilities shown in fig.

x =	-3	-2	-1	0	1	2
P(x <sub>i</sub> ) =	0.2	0.5k	k	0.1	0.3k	k

(a) Find the value of k (b) Find  $F_X(x)$  and  $f_X(x)$  also draw the plots.

Sol: we know that sum of all probabilities = 1

$$\sum P(x_i) = 1$$

$$\Rightarrow 0.2 + 0.5k + k + 0.1 + 0.3k + k = 1$$

$$0.3 + 2.8k = 1$$

$$k = \frac{0.7}{2.8} = 0.25$$

Probability function

$$F_x(x) = P\{x \leq x\}$$

$$F_x(-3) = P\{x \leq (-3)\} = 0.2$$

$$F_x(-2) = P\{x \leq (-2)\} = P\{x = -3\} + P\{x = -2\} = 0.325$$

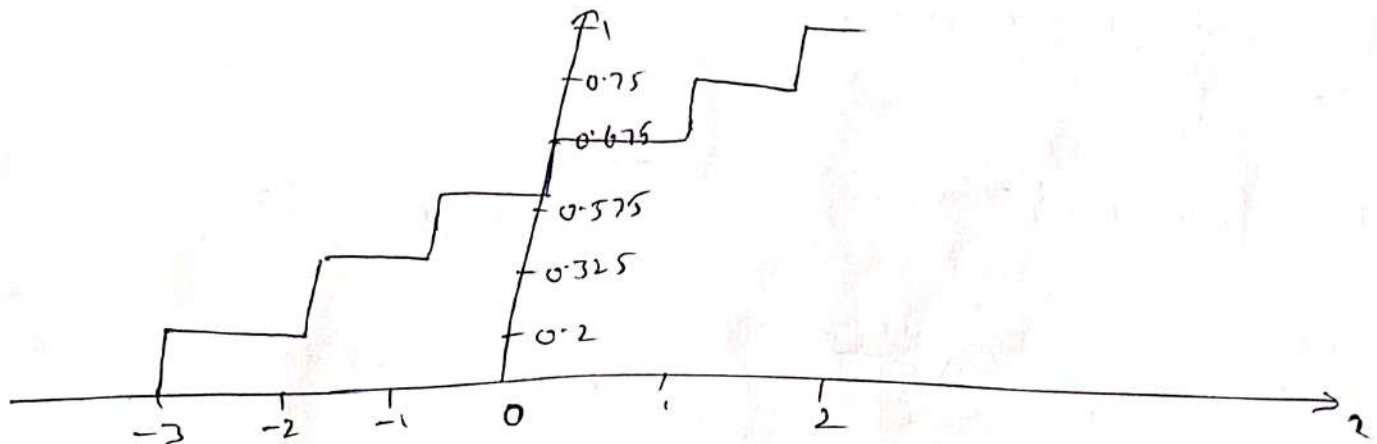
$$F_x(-1) = P\{x \leq -1\} = 0.2 + 0.5k + k = 0.575$$

$$F_x(0) = 0.575 + 0.1 = 0.675$$

$$F_x(1) = 0.675 + 0.075k = 0.75$$

$$F_x(2) = 0.75 + 0.25 = 1$$

$x_i$	-3	-2	-1	0	1	2
$P(x_i)$	0.2	0.125	0.25	0.1	0.075	0.25
$F_x(x)$	0.2	0.325	0.575	0.675	0.75	1



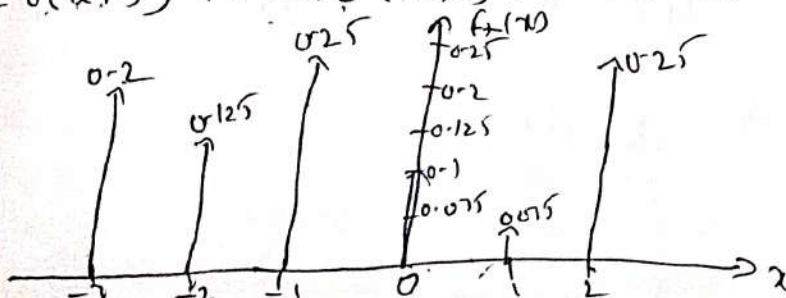
The expression of distribution function for discrete random variable  $x$  is given by  $F_x(x) = \sum_{i=-3}^2 P(x_i) u(x - x_i)$

$$F_x(x) = 0.2 u(x+3) + 0.125 u(x+2) + 0.25 u(x+1) + 0.1 u(x) + 0.075 u(x-1) + 0.25 u(x-2)$$

The expression of density function for discrete random variable is

$$f_x(x) = \sum_{i=-3}^2 P(x_i) \delta'(x - x_i)$$

$$f_x(x) = 0.2 \delta'(x+3) + 0.125 \delta'(x+2) + 0.25 \delta'(x+1) + 0.1 \delta'(x) + 0.075 \delta'(x-1) + 0.25 \delta'(x-2)$$



plot of  $f_x(x)$

→ The CDF of a random variable  $Y$

$$F_Y(y) = \begin{cases} 1 - \exp(-0.4\sqrt{y}) & y > 0 \\ 0 & y \leq 0 \end{cases} \quad \text{calculate prob of } \{2.5 < Y < 6.2\}$$

Soln:  $P\{2.5 < Y < 6.2\} = F_Y(6.2) - F_Y(2.5)$

$$F_Y(6.2) = 1 - e^{(-0.4\sqrt{6.2})} = 1 - e^{-0.4\sqrt{6.2}} \approx 0.6706$$

$$F_Y(2.5) = 1 - e^{(-0.4\sqrt{2.5})} = 0.40687$$

$$F_Y(6.2) - F_Y(2.5) = 0.1619$$

→ A random variable 'x' has the distribution function  $F_X(x) =$

$$\sum_{m=1}^{12} \frac{m^2}{650} u(x-m) \quad \text{Find the prob of (a) } P\{-\infty < x \leq 6.5\}$$

$$F_X(x) = \sum_{m=1}^{12} \frac{m^2}{650} u(x-m) \quad \text{(b) } P\{x > 4\} \quad \text{c) } P\{6 < x \leq 9\}$$

a)  $P\{-\infty < x \leq 6.5\} = F_X(6.5) - F_X(-\infty)$   
 $= F_X(6.5)$

$$F_X(6.5) = \sum_{m=1}^{12} \frac{m^2}{650} u(6.5-m)$$

$$u(6.5-m) = \begin{cases} 1 & 6.5-m \geq 0 \\ 0 & 6.5 < m \end{cases}$$

$$= \sum_{m=1}^6 \frac{m^2}{650} \times 1 = \frac{1}{650} [1^2 + 2^2 + 3^2 + 4^2 + 5^2 + 6^2] = 0.14$$

$$m=6.5$$
  

$$m=6$$

$$P\{-\infty < x \leq 6.5\} = 0.14$$

b)  $P\{x > 4\} = 1 - P\{x \leq 4\} = 1 - F_X(4)$

$$u(4-m) = \begin{cases} 1 & 4-m \geq 0 \\ 0 & \text{elsewhere} \end{cases}$$

$$F_X(4) = \sum_{m=1}^{12} \frac{m^2}{650} u(4-m)$$

$$m=4$$

$$P\{x > 4\} = 1 - \sum_{m=1}^4 \frac{m^2}{650} (1) = 1 - \frac{1}{650} \sum_{m=1}^4 m^2 = 1 - 0.046 = 0.9538$$

c)  $P\{6 < x \leq 9\} = F_X(9) - F_X(6)$

$$F_X(9) = \sum_{m=1}^{12} \frac{m^2}{650} u(9-m) = \sum_{m=1}^9 \frac{m^2}{650} (1) = 0.4384$$

$$\therefore F_X(9) - F_X(6) = 0.2984$$

→ Determine whether the following is valid distribution function (9)

$$F_x(x) = \begin{cases} 1 - e^{-x/2} & \text{for } x \geq 0 \\ 0 & \text{for } x < 0 \end{cases}$$

Soln. Given distribution function is valid if the corresponding density function is valid

$$f_x(x) = \frac{d}{dx} F_x(x) = \frac{d}{dx} (1 - e^{-x/2}) \\ = -e^{-x/2} \cdot -\frac{1}{2} = \frac{1}{2} e^{-x/2}$$

$$f_x(x) = \begin{cases} \frac{1}{2} e^{-x/2} & \text{for } x \geq 0 \\ 0 & \text{for } x < 0 \end{cases}$$

consider  $\int_{-\infty}^{\infty} f_x(x) dx = \int_{-\infty}^0 0 dx + \int_0^{\infty} \frac{1}{2} e^{-x/2} dx$

$$= \frac{1}{2} \cdot \frac{e^{-x/2}}{-1/2} \Big|_0^{\infty} = - [e^{-\infty} - e^0] = 1$$

Hence density function is valid and also distribution function is valid.

→ Let  $x$  be a continuous random variable with density function

$$f_x(x) = \begin{cases} \frac{x}{a} + k & \text{for } 0 \leq x \leq b \\ 0 & \text{elsewhere} \end{cases} \quad \text{Find (i) value of } k \quad \text{(ii) } P\{2 \leq x \leq 5\}$$

Soln. consider  $\int_{-\infty}^{\infty} f_x(x) \cdot dx = 1 = \int_0^b (\frac{x}{a} + k) dx = 1 = \frac{x^2}{2a} + kx \Big|_0^b = 1$

$$2 + 6k = 1 \Rightarrow k = -\frac{1}{6}$$

(ii)  $P\{2 \leq x \leq 5\} = \int_2^5 f_x(x) dx = \int_2^5 (\frac{x}{a} - \frac{1}{6}) dx = 0.666$

→ If the prob density of random variable is given by

$$f_x(x) = \begin{cases} c \exp(-x/4) & 0 \leq x \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

Find the value 'c' must have and also evaluate  $F_x(0.5)$ .

Soln. Given  $f_x(x) = \begin{cases} c \cdot \exp(-x/4) & 0 \leq x \leq 1 \\ 0 & \text{elsewhere} \end{cases}$

$$\int_0^1 f_x(x) \cdot dx = 1 \Rightarrow \int_0^1 c \cdot e^{-x/4} dx = 1 \Rightarrow \frac{c \cdot e^{-x/4}}{-1/4} \Big|_0^1 = 1$$

$$-4c (e^{-1/4} - 1) = 1 \Rightarrow c = \frac{1}{4(1 - e^{-0.25})} = 1.17$$

(ii)  $F_x(0.5) = ?$

$$F_x(x) = \int_{-\infty}^x f_x(x) \cdot dx$$

$$F_x(0.5) = \int_{-\infty}^{0.5} 1.13 e^{-2.14x} dx \Rightarrow 1.13 \frac{e^{-2.14x}}{-2.14} \Big|_0^{0.5}$$

$$= -4(1.13) [e^{-0.5 \cdot 2.14} - e^0] = \underline{0.53}$$

→ Find a value for constant A such that  $f_x(1) = 0$  (2-)

or a valid prob density function.

$$= A(1-x^2) \cos \frac{\pi x}{2}$$

-1 ≤ x ≤ 1  
= 0 x > 1

Soln:- For a valid density  $\int_{-\infty}^{\infty} f_x(x) dx = 1$

$$\int_{-\infty}^{\infty} A(1-x^2) \cos \frac{\pi x}{2} dx = 1 \Rightarrow A \int_{-1}^1 \cos \frac{\pi x}{2} dx - A \int_{-1}^1 x^2 \cos \frac{\pi x}{2} dx = 1$$

$$\Rightarrow A \left[ \frac{2}{\pi} \sin \frac{\pi x}{2} \right]_{-1}^1 - A \left[ x^2 \frac{2}{\pi} \sin \frac{\pi x}{2} - \right.$$



→ A random variable 'x' has a PDF  $f(x) = \begin{cases} c(1-x^4) & -1 \leq x \leq 1 \\ 0 & \text{else} \end{cases}$  (10)

(i) find 'c' (ii) find  $P\{|x| \leq \frac{1}{2}\}$ .

Sol ... we know that  $\int_{-\infty}^{\infty} f(x) \cdot dx = 1 \Rightarrow \int_{-1}^1 c(1-x^4) \cdot dx = 1$

$$c \left[ x - \frac{x^5}{5} \right]_{-1}^1 = 1 \Rightarrow c = \frac{5}{8}$$

$$f(x) = \begin{cases} \frac{5}{8}(1-x^4) & -1 \leq x \leq 1 \\ 0 & \text{else} \end{cases}$$

(ii)  $P\{|x| \leq \frac{1}{2}\} = P\left\{x \leq \frac{1}{2} \text{ or } -x \leq \frac{1}{2}\right\}$   
 $= P\left\{x \leq \frac{1}{2} \text{ or } x > -\frac{1}{2}\right\}$   
 $= P\left\{-\frac{1}{2} < x < \frac{1}{2}\right\}$   
 $= \int_{-\frac{1}{2}}^{\frac{1}{2}} f(x) \cdot dx = \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{5}{8}(1-x^4) dx$   
 $= \frac{5}{8} \left[ x - \frac{x^5}{5} \right]_{-\frac{1}{2}}^{\frac{1}{2}} = 0.617$

→ The P.D.F of a random variable x is given by  $f(x) = \begin{cases} k & a \leq x \leq b \\ 0 & \text{else} \end{cases}$

where k is constant. (i) find value of k

(ii) let  $a=1, b=2$  calculate  $P\{|x| \leq c\}$  for  $c=0.5$

we know that  $\int_a^b k \cdot dx = 1 \Rightarrow k \int_a^b dx = 1 \Rightarrow k = \frac{1}{b-a}$

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{else} \end{cases}$$

(ii)  $a=1, b=2$

$$f(x) = \begin{cases} 1 & \text{for } 1 \leq x \leq 2 \\ 0 & \text{else} \end{cases}$$

$$P\{|x| \leq 0.5\} = P\{x \leq 0.5 \text{ or } x \geq -0.5\}$$

$$= P\{-0.5 \leq x \leq 0.5\}$$

$$\int_{-0.5}^{0.5} 1 dx = 0 \text{ because } f(x) \text{ does not exist}$$

→ consider a prob density function  $f(x) = a e^{-b|x|}$  where  $x$  is a random variable. whose allowable values  $-\infty$  to  $\infty$ . Find (i) CDF  $F(x)$  (ii) Relationship b/w  $a$  &  $b$  (iii) Find prob that the outcome lies b/w 1 to 2.

Sol...  $f(x) = a e^{-b|x|}$

$$f(x) = \begin{cases} a e^{-bx} & \text{for } x \geq 0 \\ a e^{bx} & \text{for } x < 0 \end{cases}$$

(i) CDF for  $x \geq 0$ :  $F(x) = \int_{-\infty}^x f(x) \cdot dx$

$$F(x) = \int_{-\infty}^0 a e^{bx} dx + \int_0^x a e^{-bx} dx$$

$$= a \cdot \left. \frac{e^{bx}}{b} \right|_{-\infty}^0 + a \left. \frac{e^{-bx}}{-b} \right|_0^x = \frac{a}{b} [1-0] - \frac{a}{b} [e^{-bx} - 1]$$

$$= \frac{a}{b} - \frac{a}{b} e^{-bx} + \frac{a}{b} = \frac{2a}{b} - \frac{a}{b} e^{-bx} = \frac{a}{b} [2 - e^{-bx}]$$

for  $x < 0$ :

$$F(x) = \int_{-\infty}^x f(x) \cdot dx = \int_{-\infty}^x a e^{bx} dx$$

$$= \frac{a}{b} e^{bx}$$

(ii) w.k.T the total area of density function is unity.

$$\int_{-\infty}^0 a e^{bx} + \int_0^{\infty} a e^{-bx} = 1 \Rightarrow \frac{a}{b} e^{bx} \Big|_{-\infty}^0 + \frac{a}{-b} e^{-bx} \Big|_0^{\infty}$$

$$= \frac{a}{b} [1] - \frac{a}{b} [0-1] = \frac{2a}{b} = 1 \Rightarrow \boxed{2a=b}$$

(iii) The prob lies b/w 1 to 2 then

$$P\{1 < x < 2\} = \int_1^2 f(x) \cdot dx = \int_1^2 a e^{-bx} dx = \frac{-a}{b} [e^{-bx}]_1^2 = \frac{a}{b} [e^{-b} - e^{-2b}]$$

→ A random variable  $x$  has the density function  $f(x) = \begin{cases} \frac{1}{4} & -2 \leq x \leq 2 \\ 0 & \text{else} \end{cases}$   
 obtain (i)  $P\{x < 1\}$  (ii)  $P\{|x| < 1\}$  (iii)  $P\{2x > 2\}$   $P\{2x > 2\} = P\{x > 1\}$

Sol (i)  $P\{x < 1\} = \int_{-\infty}^1 f(x) dx = \frac{1}{4}$

(ii)  $P\{|x| < 1\} = P\{-1 < x < 1\} = \int_{-1}^1 \frac{1}{4} dx = \frac{1}{2}$

(iii)  $P\{x > 1\} = \int_1^{\infty} \frac{1}{4} dx = \frac{1}{4}$

→ A continuous random variable 'x' has a PDF for  $f(x) = 2x^2$   $0 \leq x \leq 1$  (11)  
 Find (i) a & b such that (i)  $P\{x \leq a\} = P\{x > a\}$  (ii)  $P\{x > b\} = 0.05$

Sol...  $P\{x \leq a\} = P\{x > a\}$

$$\int_0^a 2x^2 dx = \int_a^1 2x^2 dx \Rightarrow x^3 \Big|_0^a = x^3 \Big|_a^1$$

$$a^3 = \frac{1}{2} \Rightarrow a = \sqrt[3]{\frac{1}{2}}$$

$$P\{x > b\} = \int_b^1 2x^2 dx = x^3 \Big|_b^1 = 0.05 \Rightarrow 1 - b^3 = 0.05$$

$$b^3 = 1 - 0.05 = 1 - \frac{1}{20} = \frac{19}{20}$$

$$b = \left(\frac{19}{20}\right)^{1/3}$$

→ A random variable x has a density function  $f(x) = k \cdot \frac{1}{1+x^2}$   
 determine the value of k and also obtain distribution function

Sol w.k.t  $\int_{-\infty}^{\infty} k \cdot \frac{1}{1+x^2} dx = 1$

$$k \cdot \tan^{-1} x \Big|_{-\infty}^{\infty} = 1 \Rightarrow k \left[ \frac{\pi}{2} + \frac{\pi}{2} \right] = 1$$

$$f(x) = \frac{1}{\pi} \cdot \frac{1}{1+x^2} \quad k = \frac{1}{\pi}$$

$$F(x) = \int_{-\infty}^x f(x) \cdot dx = \int_{-\infty}^x \frac{1}{\pi} \cdot \frac{1}{1+x^2} dx$$

$$= \frac{1}{\pi} \tan^{-1} \Big|_{-\infty}^x = \frac{1}{\pi} \left[ \tan^{-1} x + \frac{\pi}{2} \right]$$

→ If the random variable x that can assume any value b/w  $x=2$  &  $x=5$  has a density function  $f(x) = k(1+x)$  find  $P\{x < 4\}$

Sol  $f(x) = \begin{cases} k(1+x) & \text{for } 2 \leq x \leq 5 \\ 0 & \text{else} \end{cases}$

$$\int_2^5 f(x) \cdot dx = 1 \Rightarrow k \int_2^5 (1+x) dx = 1 \Rightarrow k \left[ x + \frac{x^2}{2} \right]_2^5 = 1$$

$$k \left[ 1 + \frac{25}{2} \right] = 1 \Rightarrow k = \frac{2}{27}$$

$$(ii) P\{x < 4\} = \int_2^4 f(x) \cdot dx$$

$$= \int_2^4 \frac{2}{27} (1+x) dx = \frac{2}{27} \left[ x + \frac{x^2}{2} \right]_2^4$$

$$= \frac{16}{27}$$

## Examples of density and distribution functions:

The general " " " " are

- Gaussian density function (or) normal density function
- uniform " "
- Exponential " "
- Rayleigh " "
- Binomial " "
- Poissons " "

NOTE: In the above first four are continuous random variable and last two are discrete random variable.

## Gaussian density function:-

A random variable  $x$  is called gaussian then its gaussian prob density function of a random variable is given by

$$f_x(x) = \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} \quad \text{for all } 'x' - (1)$$

Here  $\sigma_x > 0$  and  $-\infty < \mu_x < \infty$  are real constants.

$\sigma_x$  = standard deviation

$\mu_x$  = Mean of the random variable

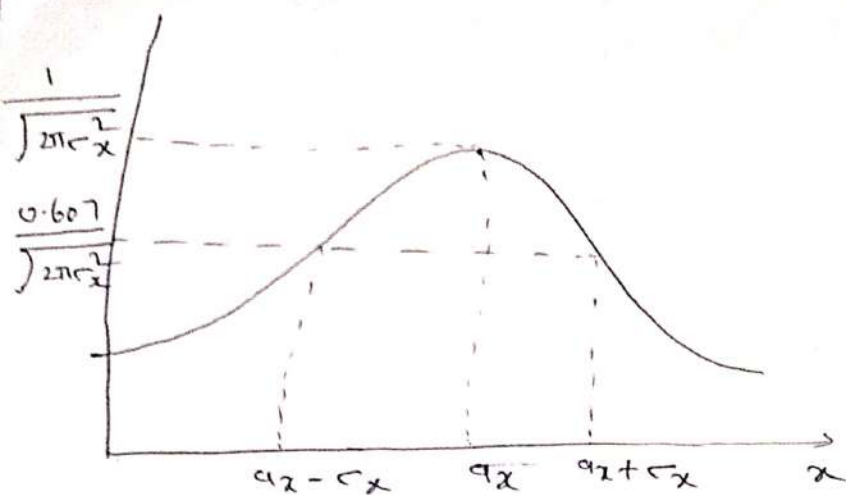
$\sigma_x^2$  = Variance of the " "

The gaussian probability distribution function is  $F_x(x) = \int_{-\infty}^x f_x(x) dx$

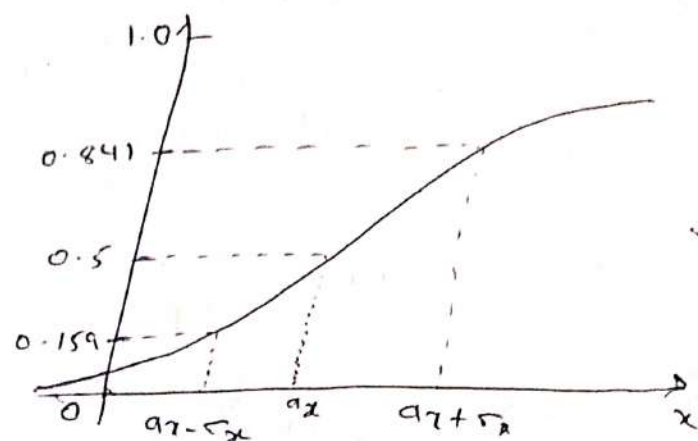
$$F_x(x) = \frac{1}{\sqrt{2\pi}\sigma_x} \int_{-\infty}^x e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} dx - (2)$$

This integral has no known closed form solution and must be evaluated by using numerical or approximation method.

We could develop a set of tables of  $F_x(x)$  for various  $x$  with  $\mu_x$  and  $\sigma_x$  parameters. This approach has limited value because there is an infinite no. of possible combinations of  $\mu_x$  and  $\sigma_x$  which requires an infinite no. of tables.



(a) Gaussian density function



FIS(b) Gaussian distribution function.

A random variable which satisfies the gaussian random variable density function is called gaussian random variable.

From fig(a)

The gaussian density function is in bell shape and is symmetrical about the mean. The maximum value of  $f_x(x)$  is  $\frac{1}{\sqrt{2\pi\sigma^2}}$  occurs at  $x = \mu$ .

The function decreases to 0.607 times of its maximum value at  $x = \mu - \sigma$  and  $x = \mu + \sigma$ .

→ From fig(b), The distribution function  $F_x(x)$  is 0.5 at  $x = \mu$ , 0.15 at  $x = \mu - \sigma$  and 0.841 at  $x = \mu + \sigma$ .

### Normalised Gaussian distribution function

If the gaussian distribution function has  $\mu = 0$  and  $\sigma = 1$  then it is called normalised gaussian distribution function. It is denoted by  $F(x)$  and

$$F(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-x^2/2} \quad \text{for } x \geq 0$$

For a negative value of  $x$  we use the  $F(-x) = 1 - F(x)$ .

To get the generalised distribution function  $F_x(x)$  in terms of normalised distribution function  $F(x)$  is

consider  $F_T(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$

let  $\frac{x-\mu}{\sigma} = u \Rightarrow \frac{1}{\sigma} dx = du$

Then  $F_T(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{x-\mu}{\sigma}} e^{-u^2/2} du = F(u) = F\left(\frac{x-\mu}{\sigma}\right)$

$$F_T(x) = F\left(\frac{x-\mu}{\sigma}\right)$$

NOTE: The function  $F(x)$  can be evaluated by using Q-function approximation. Then  $F(x) = 1 - Q(x)$ .

Q-function is defined as  $Q(x) = \frac{1}{\sqrt{2\pi}} \int_x^{\infty} e^{-z^2/2} dz$

The Q-function can be approximated as

$$Q(x) \approx \left[ \frac{1}{(1-a)x + a} \sqrt{x^2 + b} \right] \frac{e^{-x^2/2}}{\sqrt{2\pi}} \quad x \geq 0 \text{ where } a \text{ and } b \text{ are constants}$$

This approximation has been found to give minimum absolute relative error for any  $x \geq 0$  when  $a = 0.339$  and  $b = 5.510$

$$\therefore Q(x) \approx \left[ \frac{1}{0.661x + 0.339} \sqrt{x^2 + 5.51} \right] \frac{e^{-x^2/2}}{\sqrt{2\pi}} \text{ for } x \geq 0.$$

This approximation is said to equal the true value of  $Q(x)$  within a minimum absolute error of 0.270%, for any  $x \geq 0$ .

### Application:

The gaussian density is the most important of all densities and it enters into nearly all areas of science and engineering.

→ Especially in electronics and communication systems the distribution of noise signal either internally generated or externally generated exactly match with gaussian prob.

Hence it is possible to eliminate the noise completely by knowing its behaviour using gaussian density function.

Problem:- Find the prob of the event  $\{x \leq 5.5\}$  for a gaussian random variable having  $\mu_x = 3$  and  $\sigma_x = 2$ .

Soln:- Given that gaussian random variable and  $\mu_x = 3, \sigma_x = 2$   
 $P\{x \leq 5.5\} = F_x(5.5) = F\left[\frac{x - \mu_x}{\sigma_x}\right] = F\left[\frac{5.5 - 3}{2}\right] = 0.8944$

→ Assume that the height of the clouds above the ground at some location is a gaussian random variable  $x$  with  $\mu_x = 1830$  and  $\sigma_x = 460$  m. we find the prob that clouds will be higher than 2750 m

Soln:- Given that gaussian random variable with  $\mu_x = 1830$  m and  $\sigma_x = 460$  m. The prob of clouds will be higher than 2750 m is given by  
 $P\{x > 2750\} = 1 - P\{x \leq 2750\} = 1 - F_x(2750) = 1 - F\left[\frac{2750 - 1830}{460}\right]$   
 $= 1 - F(2)$   
 $= 1 - 0.9773 = 0.0227$ .

The normalized distribution function can also be approximated by using Q-function

$$Q(x) = \frac{1}{0.6612 + 0.539\sqrt{x^2 + 5.51}} e^{-x^2/2}$$

w.k.T  $F(x) = 1 - Q(x)$

$F(2) = 1 - Q(2)$

$= 1 - 0.02281 = 0.9772$

$P\{x \geq 2750\} = 1 - F(2) = 1 - 0.9773 = 0.0227$

→ For a gaussian random variable with  $\mu_x = 0$  and  $\sigma_x = 1$  what is the prob of  $P\{x > 2\}$ .

Soln:- Given g.r.v with  $\mu_x = 0$  and  $\sigma_x = 1$

$P\{x > 2\} = 1 - P\{x \leq 2\} = 1 - F_x(2) = 1 - F\left(\frac{2 - 0}{1}\right) = 1 - F(2)$

$= 0.0227$ .

## Uniform density function ::

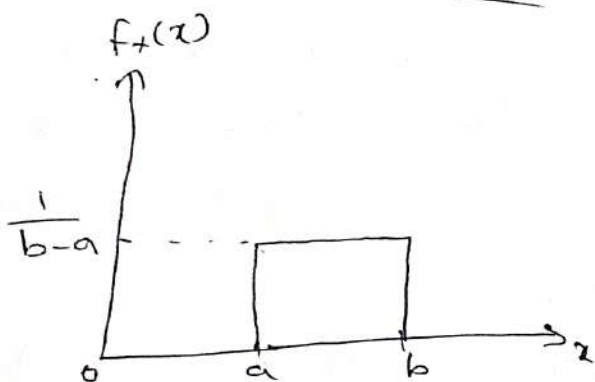
The uniform probability density function  $f(x)$  is defined as

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{for } a \leq x \leq b \\ 0 & \text{else} \end{cases} \quad \text{where } a \leq b \text{ are real constants and } -\infty < a < \infty \times b > a.$$

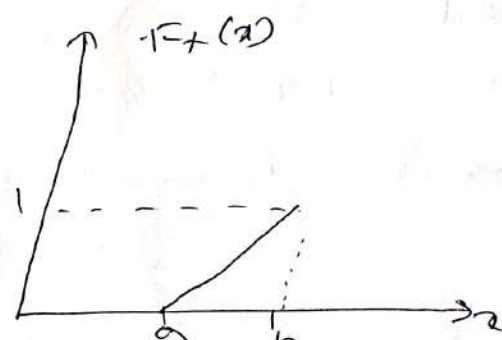
The uniform distribution function is  $F(x) = \int_{-\infty}^x f(x) \cdot dx$

$$= \int_a^x \frac{1}{b-a} dx = \frac{1}{b-a} (x-a)$$

$$F(x) = \begin{cases} 0 & \text{for } x < a \\ \frac{x-a}{b-a} & \text{for } a \leq x \leq b \\ 1 & \text{for } b \leq x \end{cases} \quad \begin{matrix} x < a \\ \\ x \geq b \end{matrix}$$



density function



uniform distribution function.

→ A particularly important application is in the quantization of signal samples prior to encoding in digital communication systems. Quantization amounts to "rounding off" the actual sample to the nearest of a large number of discrete "quantum levels". The errors introduced in the round-off process are uniformly distributed.

## Exponential density function ::

The "prob" " " is defined as

$$f(x) = \begin{cases} \frac{1}{b} e^{-(x-a)/b} & x > a \\ 0 & x < a \end{cases}$$

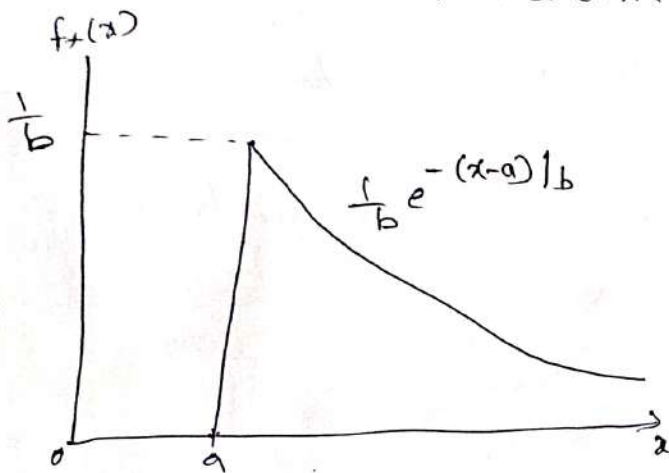


where  $a, b$  are real constants and  $-\infty < a < \infty$  and  $b > 0$

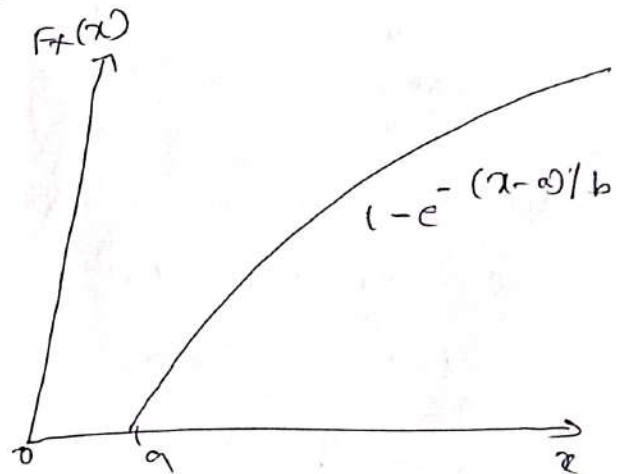
$$\begin{aligned} \text{The distribution function is } F_T(x) &= \int_{-\infty}^x f_T(x) \cdot dx \\ &= \int_a^x \frac{1}{b} e^{-(x-a)/b} dx \\ &= \frac{1}{b} \left[ \frac{e^{-(x-a)/b}}{-\frac{1}{b}} \right]_a^x \\ &= -e^{-(x-a)/b} + 1 \end{aligned}$$

$$\therefore F_T(x) = \begin{cases} 1 - e^{-(x-a)/b} & x \geq a \\ 0 & x < a \end{cases}$$

The exponential density is useful in describing raindrop sizes when a large no. of rainstorm measurements are made.



Exponential density function



distribution function

Rayleigh probability density function:-

The " " " " is defined as

$$f_T(x) = \begin{cases} \frac{2}{b} (x-a) e^{-(x-a)^2/b} & x \geq a \\ 0 & x < a \end{cases}$$

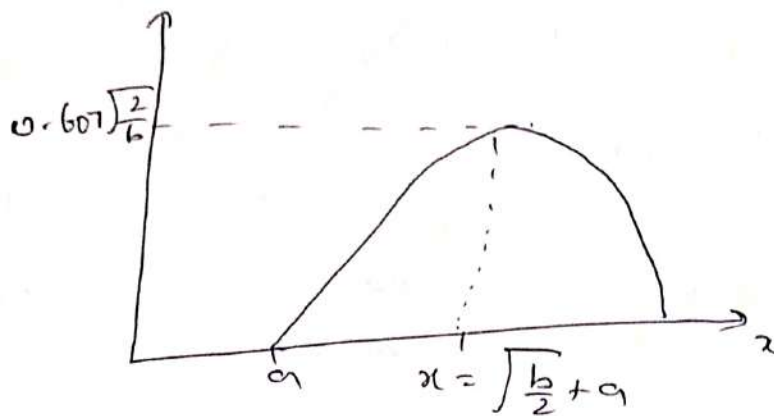
where  $a, b$  are real constants and  $-\infty < a < \infty$  and  $b > 0$

$$\begin{aligned} \text{The distribution function } F_T(x) &= \int_{-\infty}^x f_T(x) \cdot dx \\ &= \int_a^x \frac{2}{b} (x-a) e^{-(x-a)^2/b} dx \end{aligned}$$

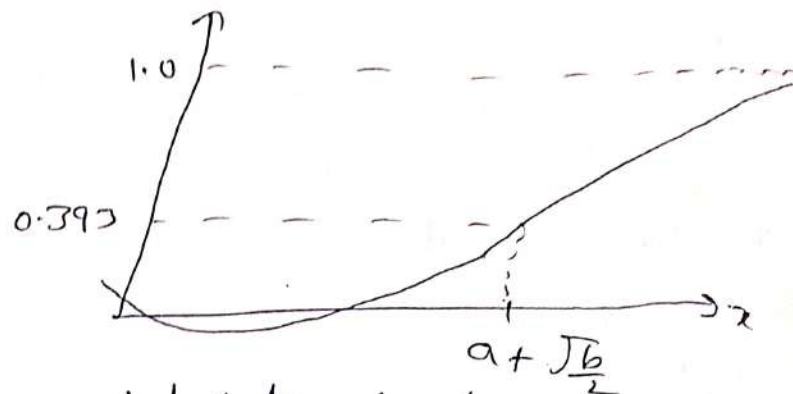
$$\text{Let } \left(\frac{x-a}{b}\right)^2 = u \Rightarrow dx \cdot \frac{2(x-a)}{b} = du$$

$$= \int_0^u e^{-u} du = \frac{e^{-u}}{-1} \Big|_0^u = e^{-u} + 1 = 1 - e^{-u} = 1 - e^{-(x-a)^2/b}$$

$$F_T(x) = \begin{cases} 1 - e^{-(x-a)^2/b} & x > a \\ 0 & x < a \end{cases}$$



Rayleigh density function



distribution function.

→ Derive the maximum value of a Rayleigh density function.

Sol: we know that the Rayleigh density function

$$f_T(x) = \frac{2}{b}(x-a) e^{-\frac{(x-a)^2}{b}} \quad \text{for } x > a$$

$$= 0 \quad \text{for } x < a$$

The maximum value occurs at  $\frac{d}{dx} f_T(x) = 0$

$$e^{-\frac{(x-a)^2}{b}} + (x-a) \left[ -e^{-\frac{(x-a)^2}{b}} \frac{d}{dx} \left( \frac{(x-a)^2}{b} \right) \right] = 0$$

$$(x-a) e^{-\frac{(x-a)^2}{b}} \cdot 2 \frac{(x-a)}{b} = e^{-\frac{(x-a)^2}{b}}$$

$$2(x-a)^2 = b$$

$$(x-a)^2 = \frac{b}{2}$$

$$x = a + \sqrt{\frac{b}{2}}$$

The maximum value of Rayleigh density function occurs at

$$x = a + \sqrt{\frac{b}{2}}$$

The maximum value of  $f_x(x)$  at  $x = a + \sqrt{\frac{b}{2}}$  is

$$= \frac{2}{b} (a + \sqrt{\frac{b}{2}} - a) \left( e^{-\frac{(a + \sqrt{\frac{b}{2}} - a)^2}{b}} \right)$$

$$= \frac{2}{b} \sqrt{\frac{b}{2}} e^{-1/2} = \sqrt{\frac{2}{b}} e^{-1/2} = 0.607 \sqrt{\frac{2}{b}}$$

Derive the maximum value of gaussian density function:

sdy w.k.T  $f_x(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-a)^2}{2\sigma^2}}$  for all  $x$

$$\frac{d}{dx} f_x(x) = 0 \Rightarrow \frac{1}{\sqrt{2\pi\sigma^2}} \frac{d}{dx} e^{-\frac{(x-a)^2}{2\sigma^2}} = 0$$

$$e^{-\frac{(x-a)^2}{2\sigma^2}} \left[ -\frac{1}{\sigma^2} \frac{d}{dx} (x-a)^2 \right] = 0$$

$$e^{-\frac{(x-a)^2}{2\sigma^2}} = 0 \quad 2(x-a) = 0$$

$$\boxed{x = a}$$

The max value is  $f_x(x) = \frac{1}{\sqrt{2\pi\sigma^2}} = \frac{1}{\sqrt{2\pi} \cdot \sigma}$

→ The Rayleigh density describes the envelope of one type of noise when passed through a bandpass filter. It also is important in analysis of errors in various measurement systems.

Binomial density function:-

Let us consider an experiment having only two possible outcomes. Let the experiment be a tossing a fair coin. ~~the prob~~ Let the falling of head corresponds to success and falling of tail corresponds to failure. Let the probability of success be "p" and the prob of failure be "q".  $[q = 1-p]$ . such that  $p+q=1$ .

The experiment is repeated "N" trials then the Binomial probability ~~function~~ density function of a discrete R.V. "x" is defined as

$$f_x(x) = \sum_{k=0}^N \binom{N}{k} p^k (1-p)^{N-k} \delta(x-k) \quad \text{--- (1)}$$

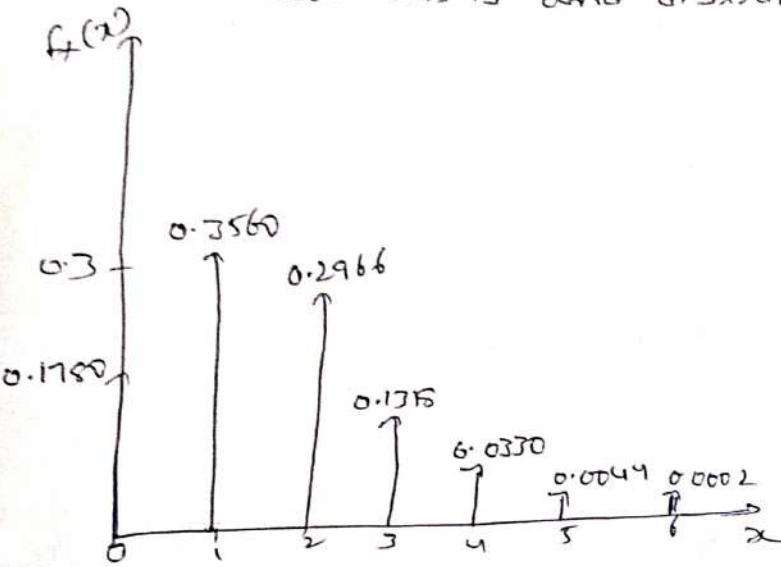
and the binomial distribution is

$$F_x(x) = \sum_{k=0}^N \binom{N}{k} p^k (1-p)^{N-k} u(x-k) \quad (2)$$

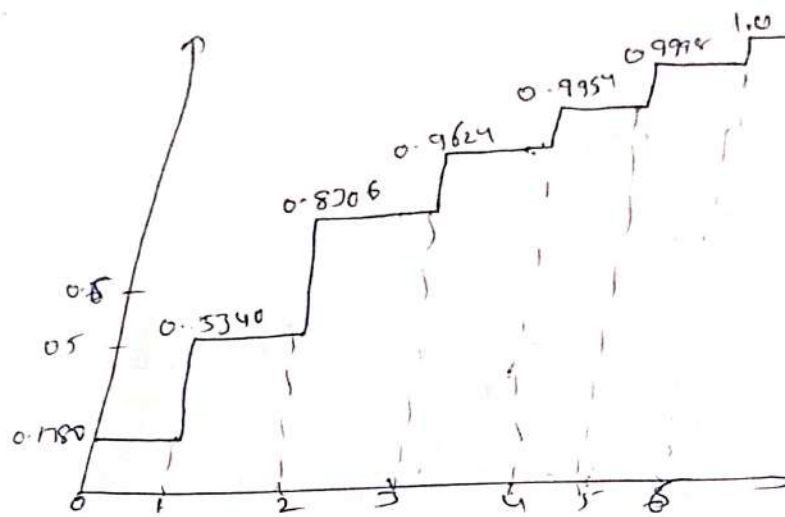
↓ This is obtained by Bernoulli's trials

where  $\binom{N}{k}$  is equal to  $\frac{N!}{k!(N-k)!}$

The binomial density and distribution function for  $N=6$  &  $p=0.25$ .



(a) density function



(b) distribution function.

### Applications:

The Binomial density function can be used in the following applications.

1. Bernoulli's trial experiment
2. It applies to many games of chances.
3. Detection problems in radar and sonar
4. Many experiments having only two possible outcomes on a given trial.

### Poisson's probability density function:

The " " " " " for a discrete random

variable  $x$  is given by  $f_x(x) = e^{-b} \sum_{k=0}^{\infty} \frac{b^k}{k!} \delta(x-k) \quad (1)$

The Poisson's prob distribution is

$$F_x(x) = e^{-b} \sum_{k=0}^{\infty} \frac{b^k}{k!} u(x-k) \quad (2)$$

The Poisson distribution is approximated function for Binomial distribution for  $n \rightarrow \infty$  and  $p \rightarrow 0$ , Here the constant  $b = np$ .

Applications:

- It is mostly applied to counting type problems.
- it describes
- The no. of Telephone calls made during a period of time
- The " " defective elements on a given samples
- The no. of electrons emitted from a cathode on a given time interval.

NOTE: If the time interval is 'T' sec and the average rate of counting the items is  $\lambda$  then in the distribution  $b$  is given by

$$b = \lambda T$$

Problem:-

Assume Automobile arrives at a gasoline station are Poisson and occurs at an average rate of 50 per hour. The station has only one gasoline pump. If the all the cars are assume to require 1 minute to obtain fuel. what is the Prob that a waiting line will occur at the pump?

Sol: Given automobile arrives at a gasoline station are Poisson  
Average rate  $\lambda = 50$  per hour, Time duration  $T = 1$  minute

$$b = \lambda T = \frac{50}{60} \times 1 = \frac{5}{6}$$

A waiting line occurs if two or more cars arrive in one minute duration.

The Prob. of waiting line equals to  $P\{K \geq 2\}$ .

$$= 1 - P\{K \leq 1\} = 1 - F_x(1)$$

w.k.T

$$F_x(x) = e^{-b} \sum_{k=0}^{\infty} \frac{b^k}{k!} u(x-k)$$

$$= e^{-\frac{5}{6}} \sum_{k=0}^{\infty} \frac{(\frac{5}{6})^k}{k!} u(1-k)$$

$$= e^{-\frac{5}{6}} \sum_{k=0}^1 \frac{(\frac{5}{6})^k}{k!} u(1-k)$$

$$= e^{-\frac{5}{6}} \cdot \frac{(\frac{5}{6})^0}{0!} + e^{-\frac{5}{6}} \frac{(\frac{5}{6})^1}{1!} = e^{-\frac{5}{6}} + e^{-\frac{5}{6}} \cdot \frac{5}{6} = 0.7967$$

$$1 - F(x) = 0.2032$$

## Conditional distribution & density function :-

Let  $A$  &  $B$  are events the conditional probability of event  $A$  given  $B$  is  $P(A|B) = \frac{P(A \cap B)}{P(B)}$

## conditional distribution function :-

Let  $A$  is an event  $\{x \leq x\}$ , for a random variable 'x' then the conditional distribution function of  $x$ , when event  $B$  is known is denoted by  $F_x(x|B)$  & is defined as

$$F_x(x|B) = P\{x \leq x|B\} = \frac{P\{x \leq x \cap B\}}{P(B)} \quad \text{--- (1)}$$

Here probability of  $P\{x \leq x \cap B\}$  indicates the probability of joint event  $\{x \leq x \cap B\}$

NOTE :- For discrete random variable

$$F_x(x|B) = \sum_{i=1}^N P(x_i|B) u(x-x_i) \quad \text{--- (2)}$$

Properties :- The properties of conditional distribution function  $F_x(x|B)$  are given by

$$\rightarrow F(-\infty|B) = 0 \quad \rightarrow F_x(\infty|B) = 1 \quad \rightarrow 0 \leq F_x(x|B) \leq 1$$

$$\rightarrow F_x(x_1|B) \leq F_x(x_2|B) \text{ if } x_1 < x_2$$

$$\rightarrow P\{x_1 < x < x_2|B\} = F_x(x_2|B) - F_x(x_1|B)$$

$$\rightarrow F_x(x^+|B) = F_x(x|B)$$

NOTE :- we observe that except conditional event  $B$  all properties are similar to the probability distribution function.

conditional density function :- conditional density function of a random variable is defined as derivative of C.D.F. mathematically it can be represented as

$$f_x(x|B) = \frac{d}{dx} F_x(x|B) \quad \text{--- (3)}$$

For discrete random variable  $X$

$$f_x(x|D) = \sum_{i=1}^N P(x_i|D) \delta(x-x_i) \quad \text{--- (2)}$$

Properties: conditional density function  $f_x(x|D)$  has the following properties

1.  $f_x(x|D) \geq 0$  (non-negative quantity)
2.  $\int_{-\infty}^{\infty} f_x(x|D) dx = 1$
3.  $F_x(x|D) = \int_{-\infty}^x f_x(x|D) dx$
4.  $P\left\{x_1 < x < \frac{x_2}{A}\right\} = \int_{x_1}^{x_2} f_x(x|D) dx$ .

Problem: In an experiment there are 2 boxes each box contains balls as shown in table. The event is to select a box randomly and a ball from the selected box. The probability of selecting the first box is 0.3. Also find out (i) the conditional prob distribution & density function (ii) the prob distribution & density function (iii) plot the graphs for the balls

Ball colour (x)	Boxes		Total
	1 (B <sub>1</sub> )	2 (B <sub>2</sub> )	
Red (x <sub>1</sub> )	10	50	60
Blue (x <sub>2</sub> )	20	40	60
White (x <sub>3</sub> )	30	120	150
	110	210	

Sol: Let  $D_0 x_1$  is  $B_1$ ,  $D_0 x_2$  is  $B_2$

given  $P(B_1) = 0.3$

we know that  $P(B_1) + P(B_2) = 1$   
 $P(B_2) = 0.7$

probability of selecting a red ball from  $B_1$  balls  $P(x_1|B_1) = \frac{10}{110} = \frac{1}{11}$   
 " " " " " " "  $P(x_2|B_1) = \frac{20}{110} = \frac{2}{11}$   
 " " " " " " "  $P(x_3|B_1) = \frac{30}{110} = \frac{3}{11}$   
 " " " " "  $B_2$  balls  $P(x_1|B_2) = \frac{50}{210} = \frac{5}{21}$   
 " " " " " " "  $P(x_2|B_2) = \frac{40}{210} = \frac{4}{21}$   
 " " " " " " "  $P(x_3|B_2) = \frac{120}{210} = \frac{4}{7}$

now the conditional distribution defined as

$$F_+(x|D_1) = \sum_{i=1}^3 P(x_i|D_1) u(x-x_i) = \sum_{i=1}^3 P(x_i|D_1) u(x-x_i)$$

$$= P(x_1|D_1) u(x-1) + P(x_2|D_1) u(x-2) + P(x_3|D_1) u(x-3)$$

$$= \frac{1}{11} u(x-1) + \frac{2}{11} u(x-2) + \frac{8}{11} u(x-3)$$

$$F_+(x|D_2) = \sum_{i=1}^3 P(x_i|D_2) u(x-x_i) = \sum_{i=1}^3 P(x_i|D_2) u(x-x_i)$$

$$= P(x_1|D_2) u(x-1) + P(x_2|D_2) u(x-2) + P(x_3|D_2) u(x-3)$$

$$= \frac{5}{21} u(x-1) + \frac{4}{21} u(x-2) + \frac{12}{21} u(x-3)$$

$$f_+(x|D_1) = \sum_{i=1}^3 P(x_i|D_1) \delta(x-x_i)$$

$$= P(x_1|D_1) \delta(x-1) + P(x_2|D_1) \delta(x-2) + P(x_3|D_1) \delta(x-3)$$

$$= \frac{1}{11} \delta(x-1) + \frac{2}{11} \delta(x-2) + \frac{8}{11} \delta(x-3)$$

$$f_+(x_i|D_2) = P(x_1|D_2) \delta(x-1) + P(x_2|D_2) \delta(x-2) + P(x_3|D_2) \delta(x-3)$$

$$= \frac{5}{21} \delta(x-1) + \frac{4}{21} \delta(x-2) + \frac{12}{21} \delta(x-3)$$

Probability distribution function is

$$F_+(x) = \sum_{i=1}^3 P(x_i) u(x-x_i)$$

$$= P(x_1) u(x-1) + P(x_2) u(x-2) + P(x_3) u(x-3)$$

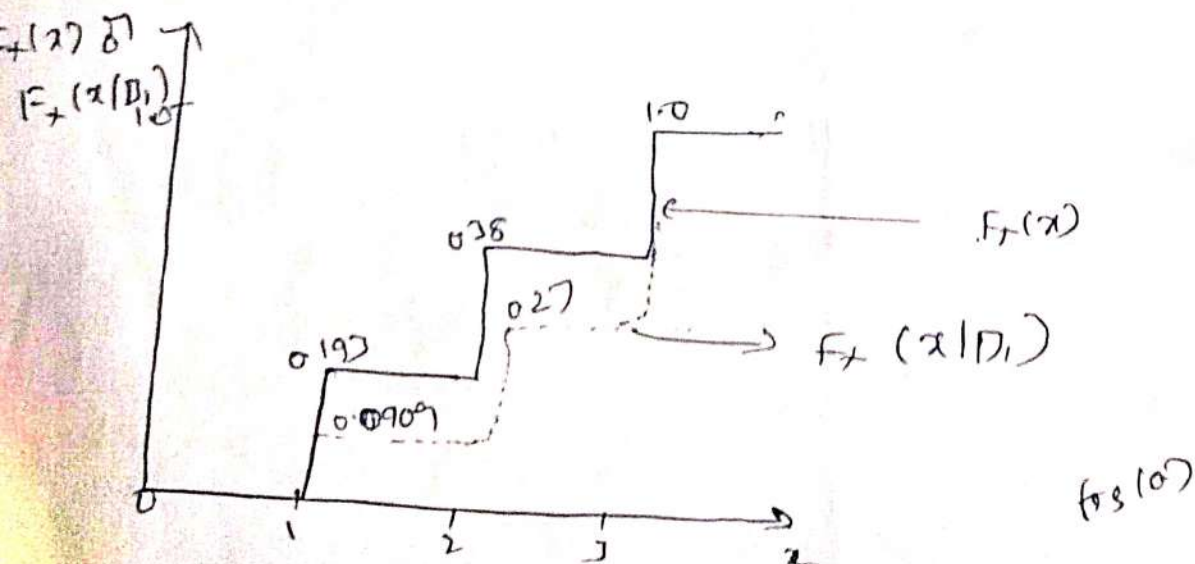
$$P(x_1) = P(x_1|D_1) P(D_1) + P(x_1|D_2) P(D_2) = \frac{1}{11} (0.3) + \frac{5}{21} (0.7) = 0.193$$

$$P(x_2) = P(x_2|D_1) P(D_1) + P(x_2|D_2) P(D_2) = \frac{2}{11} (0.3) + \frac{4}{21} (0.7) = 0.187$$

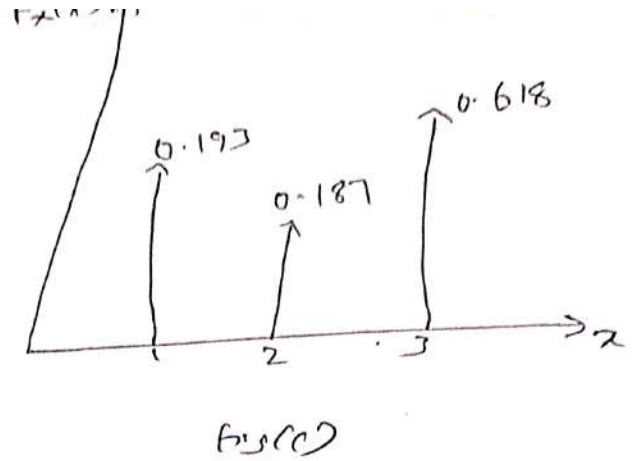
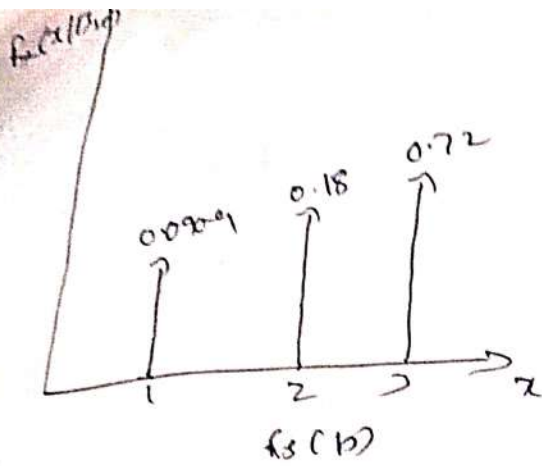
$$P(x_3) = P(x_3|D_1) P(D_1) + P(x_3|D_2) P(D_2) = \frac{8}{11} (0.3) + \frac{12}{21} (0.7) = 0.618$$

$$F_+(x) = 0.193 u(x-1) + 0.187 u(x-2) + 0.618 u(x-3)$$

$$f_+(x) = 0.193 \delta(x-1) + 0.187 \delta(x-2) + 0.618 \delta(x-3)$$







distributions (a) and densities (b) and (c).

→ Life time IC chips manufactured by a semiconductor manufacturer is approximately normally distributed with mean  $5 \times 10^6$  hours and standard deviation  $5 \times 10^5$  hours. A main frame manufacturer requires that at least 95% of a batch should have a lifetime greater than  $4 \times 10^6$  hours with the deal to make.

soln:- given  $\mu = 5 \times 10^6$  hours,  $\sigma = 5 \times 10^5$  which is a gaussian distributed function

The prob. of getting lifetime greater than  $4 \times 10^6$  hours is given by  $= P\{x > 4 \times 10^6\} = 1 - P\{x \leq 4 \times 10^6\} = 1 - F_x(4 \times 10^6)$ .

$$F_x(4 \times 10^6) = F\left(\frac{x - \mu}{\sigma}\right) = F\left[\frac{4 \times 10^6 - 5 \times 10^6}{5 \times 10^5}\right] = F\left(-\frac{10}{5}\right) = F(-2)$$